

Multiple Imputation

VGSM, Chapter 11

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Last Week

- Discussed developing models for imputing
- Look at Stata's capabilities
- Implemented imputations in Stata

UNOS Dataset

- Data on 9775 pediatric kidney transplants from 1990-2002
- Followed for transplant outcomes
465 deaths
- Predictors: age, age of donor, year, HLA loci, txtype
- Cold ischemia time: missing in 23%

Multiple Imputation

1. Assume data MAR
2. Develop predictive model for missing data
3. Sample from predictive model
4. Make multiple complete datasets
5. Analyze datasets
6. Synthesize results

Chained Syntax

- `mi register imputed hlamat age_don age cold_isc prevtx death txtype`
register everything
- `mi impute chained (pmm) cold_isc (ologit) hlamat (regress) age_don (regress) age (logit) prevtx = death txtype, add(50) rseed(897)`
- *(method) variable = nonmissing*
- Very flexible syntax

Chained Output

```
. mi impute chained (pmm) cold_isc (ologit) hlamat (regress) age_don (regress) age (logit)
prevtx = death txttype, add(20) rseed(897) dots
```

Conditional models:

```
    age: regress age i.prevtx age_don i.hlamat cold_isc death txttype
  prevtx: logit prevtx age age_don i.hlamat cold_isc death txttype
 age_don: regress age_don age i.prevtx i.hlamat cold_isc death txttype
   hlamat: ologit hlamat age i.prevtx age_don cold_isc death txttype
 cold_isc: pmm cold_isc age i.prevtx age_don i.hlamat death txttype
```

Performing chained iterations:

```
  imputing m=1 through m=20 .....10.....20 done
```

mi impute reply

```
Multivariate imputation          Imputations =      20
Chained equations                added =      20
Imputed: m=1 through m=20       updated =       0

Initialization: monotone        Iterations =     200
                                burn-in =      10
```

```
cold_isc: predictive mean matching
hlamat: ordered logistic regression
age_don: linear regression
age: linear regression
prevtx: logistic regression
```

Variable	Observations per m			Total
	Complete	Incomplete	Imputed	
cold_isc	7525	2250	2250	9775
hlamat	9541	234	234	9775
age_don	9662	113	113	9775
age	9766	9	9	9775
prevtx	9702	73	73	9775

(complete + incomplete = total; imputed is the minimum across m of the number of filled-in observations.)

MI Dataset

tabulate _mi_m

_mi_m	Freq.	Percent	Cum.
0	9,367	18.82	18.82
1	2,020	4.06	22.88
2	2,020	4.06	26.94
3	2,020	4.06	31.00
4	2,020	4.06	35.06
5	2,020	4.06	39.12
6	2,020	4.06	43.18
7	2,020	4.06	47.23
8	2,020	4.06	51.29
9	2,020	4.06	55.35
10	2,020	4.06	59.41
11	2,020	4.06	63.47
12	2,020	4.06	67.53
13	2,020	4.06	71.59
14	2,020	4.06	75.65
15	2,020	4.06	79.71
16	2,020	4.06	83.76
17	2,020	4.06	87.82
18	2,020	4.06	91.88
19	2,020	4.06	95.94
20	2,020	4.06	100.00
Total	49,767	100.00	

← original data

← imputation "10"

9,347 record dataset
becomes 49,767
records

Multiple Imputation

1. Assume data MAR
2. Develop predictive model for missing data
3. Sample from predictive model
4. Make multiple complete datasets
5. Analyze datasets
6. Synthesize results

Source of Uncertainty

- Predictive distn of cold_isc
 - $\text{cold_isc}_i = \alpha + \beta_1 * x_1 + \beta_2 * x_2 + \dots \beta_p * x_p + \epsilon_i$
- Sources of uncertainty
 1. prediction: ϵ_i
 2. parameters: $\beta_1, \beta_2, \dots, \beta_p$
 3. form of model: *predictors, transformation*
- Imputation deals with 1 & 2

Result of 1st Imputation

Logistic regression

Number of obs = 9367

LR chi2(12) = 168.53

Prob > chi2 = 0.0000

Log likelihood = -1633.2037

Pseudo R2 = 0.0491

death	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
cold_isc	1.000987	.0065701	0.15	0.881	.9881921	1.013947
age_don	.9979897	.0039913	-0.50	0.615	.9901974	1.005843
age	.9584817	.0090968	-4.47	0.000	.9408171	.9764779
year	.8574911	.0130516	-10.10	0.000	.8322883	.8834571
prevtx	1.109361	.14694	0.78	0.433	.8557115	1.438197
txtype	1.248548	.2388952	1.16	0.246	.8580995	1.816656
hlamat						
1	1.004009	.199318	0.02	0.984	.6803856	1.481562
2	.818982	.1662035	-0.98	0.325	.5502147	1.219036
3	.6860487	.1422434	-1.82	0.069	.4569506	1.030008
4	.6296867	.1537688	-1.89	0.058	.3901773	1.016218
5	.8229399	.2596461	-0.62	0.537	.4434097	1.527323
6	.8010403	.2694053	-0.66	0.509	.414361	1.548566

Imputations

Imputation #	OR txttype	SE(OR)
1	1.25	0.24
2	1.13	0.22
3	1.21	0.23
4	1.29	0.25
5	1.24	0.24

mean

1.22

0.24

SD

0.06

Combining Imputations

$K = \text{Number of Imputations} = 5$

$$\begin{aligned}\text{MI Estimate} &= (OR_1 + OR_2 + \dots + OR_K)/K \\ &= 1.22\end{aligned}$$

$$\begin{aligned}\text{MI Variance} &= (1 + 1/K)*\{ SE(OR_k) \}^2 + \\ &\quad \{ \text{mean}(\{SE_k(OR)\}^2) \} \\ &= (1 + 1/5)*\{ 0.06 \}^2 + \\ &\quad (0.24)^2 \\ &= .062\end{aligned}$$

$$\text{MI SE} = \text{sqrt}(\text{MI Variance}) = 0.25$$

Imputed Regression

the analysis step

- `mi estimate, or : logistic death cold_isc
txty i.hlam age_don age year`
- Fits a logistic regression with imputations based on the last commands
- After “:” usual commands
- Need “or” to get odds ratios

Results

```
mi estimate, or: logistic death cold_isc age_don age year prevtx txttype i.hlamat
```

```
Multiple-imputation estimates      Imputations      =      5
Logistic regression               Number of obs     =     9367
```

death	exp(b)	Std. Err.	t	P> t	[95% Conf. Interval]	
cold_isc	1.002313	.0071502	0.32	0.746	.9883722	1.01645
age_don	.9979795	.0039896	-0.51	0.613	.9901907	1.00583
age	.9584372	.0090968	-4.47	0.000	.9407726	.9764335
year	.8578479	.0130779	-10.06	0.000	.8325948	.8838669
prevtx	1.108886	.1468931	0.78	0.435	.855321	1.437622
txttype	1.216516	.2422243	0.98	0.325	.8233093	1.797515
hlamat						
1	1.002488	.1990982	0.01	0.990	.679246	1.479555
2	.8181683	.1660688	-0.99	0.323	.549628	1.217914
3	.6849142	.1420574	-1.82	0.068	.4561307	1.02845
4	.6284968	.1535163	-1.90	0.057	.3893938	1.014419
5	.8206398	.2589599	-0.63	0.531	.4421287	1.523199
6	.7986734	.26867	-0.67	0.504	.4130751	1.544221
_cons	6.3e+131	1.9e+133	9.98	0.000	8.0e+105	4.9e+157

PMM Results

```
mi estimate, or: logistic death cold_isc age_don age year prevtx txttype i.hlamat
```

Multiple-imputation estimates	Imputations	=	5
Logistic regression	Number of obs	=	9367

death	exp(b)	Std. Err.	t	P> t	[95% Conf. Interval]	
cold_isc	1.002321	.0070415	0.33	0.741	.9886026	1.01623
age_don	.9979781	.0039895	-0.51	0.613	.9901895	1.005828
age	.9584332	.0090967	-4.47	0.000	.9407689	.9764293
year	.8577968	.0130501	-10.08	0.000	.8325966	.8837597
prevtx	1.108817	.1468895	0.78	0.436	.8552597	1.437547
txttype	1.216574	.2400089	0.99	0.320	.8263761	1.791016
hlamat						
1	1.002359	.1990838	0.01	0.991	.6791438	1.479398
2	.8180324	.1660531	-0.99	0.322	.5495211	1.217746
3	.6847243	.1420257	-1.83	0.068	.4559942	1.028187
4	.6284921	.1535049	-1.90	0.057	.3894033	1.014379
5	.8208394	.2589845	-0.63	0.531	.4422767	1.523429
6	.798598	.2686421	-0.67	0.504	.4130386	1.544066
_cons	7.1e+131	2.1e+133	10.00	0.000	1.0e+106	4.9e+157

Data Management

- Imputing creates phantom observations
- Data is only good for imputations
- Useful command -- “mi extract”
creates a dataset based on one
“imputation”
- *mi extract 0, clear*
back to regular incomplete data
- *mi extract 10, clear*
complete dataset using imputation #10
allows you to check imputation set

Imputation 10

```
. logistic death cold_isc age_don age year prevtx txttype i.hlamat
```

```
Logistic regression                                Number of obs   =      9367
                                                    LR chi2(12)      =      168.51
                                                    Prob > chi2       =      0.0000
Log likelihood = -1633.2127                      Pseudo R2        =      0.0491
```

death	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
cold_isc	1.000447	.0065814	0.07	0.946	.9876305	1.01343
age_don	.997997	.0039921	-0.50	0.616	.9902032	1.005852
age	.9584903	.0090981	-4.47	0.000	.9408232	.9764891
year	.8573443	.0130516	-10.11	0.000	.8321415	.8833103
prevtx	1.109332	.1469424	0.78	0.433	.855679	1.438175
txttype	1.261713	.2416046	1.21	0.225	.8668913	1.836355
hlamat						
1	1.004259	.1993966	0.02	0.983	.6805166	1.482015
2	.8191515	.1662461	-0.98	0.326	.5503179	1.219312
3	.6863007	.1423103	-1.82	0.069	.4570993	1.03043
4	.630004	.1538591	-1.89	0.059	.3903584	1.016771
5	.8238943	.259907	-0.61	0.539	.4439664	1.528949
6	.8018133	.2697156	-0.66	0.511	.4147099	1.550251
_cons	2.0e+132	6.1e+133	10.03	0.000	2.8e+106	1.4e+158

but this dropped all₁₈ other imputations

mi xeq

simpler approach

- mi xeq l0: summ cold_isc
summarize cold_isc from imputation l0

```
m=10 data:  
-> summ cold_isc
```

Variable	Obs	Mean	Std. Dev.	Min	Max
cold_isc	9367	10.04335	11.53007	-20.96038	72

mi xeq

examining results

```
. mi xeq 10: logistic death cold_isc age_don age year prevtx txttype i.hlamat
```

m=10 data:

```
-> logistic death cold_isc age_don age year prevtx txttype i.hlamat
```

Logistic regression	Number of obs	=	9367
	LR chi2(12)	=	168.51
	Prob > chi2	=	0.0000
Log likelihood = -1633.2127	Pseudo R2	=	0.0491

death	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
cold_isc	1.000447	.0065814	0.07	0.946	.9876305	1.01343
age_don	.997997	.0039921	-0.50	0.616	.9902032	1.005852
age	.9584903	.0090981	-4.47	0.000	.9408232	.9764891
year	.8573443	.0130516	-10.11	0.000	.8321415	.8833103
prevtx	1.109332	.1469424	0.78	0.433	.855679	1.438175
txttype	1.261713	.2416046	1.21	0.225	.8668913	1.836355
hlamat						
1	1.004259	.1993966	0.02	0.983	.6805166	1.482015
2	.8191515	.1662461	-0.98	0.326	.5503179	1.219312
3	.6863007	.1423103	-1.82	0.069	.4570993	1.03043
4	.630004	.1538591	-1.89	0.059	.3903584	1.016771
5	.8238943	.259907	-0.61	0.539	.4439664	1.528949
6	.8018133	.2697156	-0.66	0.511	.4147099	1.550251
_cons	2.0e+132	6.1e+133	10.03	0.000	2.8e+106	1.4e+158

mi extract

- Useful command
- Analysis not supported in mi estimate:
- Use “mi extract, `k`” looped for k
- Save results
- Analyze using published formulae
not optimal, but possible

How Many Imputations

- Feasibly many:
“50% missingness 5 draws is highly efficient”
- True but gives poor standard errors
better guideline 20-40
- Stata on UNOS data
 - 20 imputations in 6 sec.
 - 50 imputations in 16 sec.
 - 500 imputations in 360 sec.

50 Imputations

```
. mi estimate, or: logistic death cold_isc age_don age year prevtx txttype i.hlamat
```

Multiple-imputation estimates	Imputations	=	50
Logistic regression	Number of obs	=	9367
	Average RVI	=	0.0126
DF adjustment: Large sample	DF: min	=	2492.02
	avg	=	2.42e+10
	max	=	2.69e+11
Model F test: Equal FMI	F(12, 3.2e+06)	=	13.04
Within VCE type: OIM	Prob > F	=	0.0000

death	exp(b)	Std. Err.	t	P> t	[95% Conf. Interval]	
cold_isc	1.001093	.0071376	0.15	0.878	.9871925	1.015189
age_don	.9979925	.003991	-0.50	0.615	.9902007	1.005846
age	.9584724	.0090979	-4.47	0.000	.9408057	.9764708
year	.8575193	.0130689	-10.09	0.000	.8322835	.8835204
prevtx	1.109199	.14693	0.78	0.434	.8555691	1.438017
txttype	1.245769	.2475528	1.11	0.269	.8438537	1.839111
hlamat						
1	1.003516	.1992915	0.02	0.986	.6799574	1.481042
2	.8187808	.1661875	-0.99	0.325	.5500469	1.218809
3	.6857152	.1422276	-1.82	0.069	.4566588	1.029664
4	.6294577	.153747	-1.90	0.058	.389994	1.015957
5	.822776	.2596321	-0.62	0.536	.4432816	1.527156
6	.8006779	.2693564	-0.66	0.509	.4140995	1.548143

500 Imputations

Multiple-imputation estimates	Imputations	=	500
Logistic regression	Number of obs	=	9367
	Average RVI	=	0.0130
DF adjustment: Large sample	DF: min	=	23812.19
	avg	=	1.54e+11
	max	=	1.54e+12
Model F test: Equal FMI	F(12, 3.1e+07)	=	13.02
Within VCE type: OIM	Prob > F	=	0.0000

death	exp(b)	Std. Err.	t	P> t	[95% Conf. Interval]	
cold_isc	1.002006	.0070717	0.28	0.776	.9882409	1.015964
age_don	.9979834	.0039899	-0.50	0.614	.9901938	1.005834
age	.9584474	.009097	-4.47	0.000	.9407824	.9764441
year	.8577653	.0130703	-10.07	0.000	.8325267	.8837691
prevtx	1.109026	.1469091	0.78	0.435	.8554327	1.437797
txtype	1.223848	.2423487	1.02	0.308	.8301736	1.804205
hlamat						
1	1.002682	.1991348	0.01	0.989	.6793796	1.479836
2	.8183211	.1660985	-0.99	0.323	.5497324	1.218137
3	.6850761	.1420932	-1.82	0.068	.4562356	1.028699
4	.6287796	.1535788	-1.90	0.057	.3895769	1.014854
5	.8212404	.259145	-0.62	0.533	.4424569	1.524297
6	.7992063	.2688505	-0.67	0.505	.4133495	1.545256

Approach to Imputation Model

- Inclusivity over parsimony
more predictor: MAR more plausible
- Will remove biases -- primary objective
- Two things to include: outcome, predictor in your analytic model
- Functional form tends not to be critical

Making Up Data?

- No.
- Making a model for missingness
- Making guess at missing data
- Fully representing uncertainty
 - this is the key to it's validity

Inverse Weighting for Missing Data

UNOS Dataset

- Data on 9775 pediatric kidney transplants from 1990-2002
- Followed for transplant outcomes
465 deaths
- Predictors: age, age of donor (age_don), year, # HLA loci (hlatmat), living v. cadaveric donor (txtype) and cold ischemia time
- Cold ischemia time: missing in 23%
- Ignore missingness in other variables

MAR Assumption

- Cold Ischemia values MAR with respect to some covariates and outcome (C_{MAR})
same assumption as in multiple imputation
- R: (1=cold ischemia observed, 0=missing)
- Formally
$$\frac{\text{pr}(R=1 | X_{\text{age}}, \dots, X_{\text{cold_isc}}, \dots, X_{\text{hlamat}}, Y)}{\text{pr}(R=1 | X_{\text{age}}, \dots, X_{\text{hlamat}}, Y)} = 1$$
- Probability of being observed doesn't depend on $X_{\text{cold_isc}}$
- Depends on $C_{MAR}(X_{\text{age}}, \dots, X_{\text{hlamat}}, Y)$
like a “stratified sample”

Implication #2

- Inverse weighting uses “Implication #2”

$$\text{pr}(R=1 | X_{\text{age}}, \dots, X_{\text{cold_isc}}, \dots, X_{\text{hlamat}}, Y) =$$

$$\text{pr}(R=1 | X_{\text{age}}, \dots, X_{\text{hlamat}}, Y)$$

doesn't depend on $X_{\text{cold_isc}}$

- Denote $C_{\text{MAR}} = \{X_{\text{age}}, \dots, X_{\text{hlamat}}, Y\}$

the set of MAR variables

- $\text{Pr}(R_i=1 | C_{\text{MAR}}) = \pi_i$

- Behave like a probability sample
with unknown probability of selection
But it is estimable

Steps :Inverse Weighting

1. Assume Data is Missing at Random (MAR)
2. Specify Model for Complete Data
3. Estimate/Calculate Weights
4. Analyze Available Data with Weights

MAR Assumption

Step I

- Cold Ischemia values MAR with respect to other model covariates and outcome
same assumption as in multiple imputation
- Denote $C_{\text{MAR}} = \{X_{\text{age}}, \dots, X_{\text{hlatmat}}, Y\}$
as the set of MAR variables
- Want to include Y
- Should be inclusive: more variables
similar consideration to imputation
- But “Congeniality” not a consideration

UNOS Analysis Model

Step 2

Model for the complete data

- Y: survival after transplant
- X: age, age_don, year, hlamat, txtype, cold_isc
- $$\text{logit}(Y) = \beta_0 + \beta_{\text{age}} * \text{age} + \beta_{\text{age_don}} * \text{age_don} + \beta_{\text{year}} * \text{year} + \beta_{\text{txtype}} * \text{txtype} + \beta_{\text{cold_isc}} * \text{cold_isc} + \beta_{\text{hlamat}} * \text{hlamat}$$
- Just a logistic regression

Model for Weights

Step 3

Model for probability of observed data

- Probability being observed not known
must be estimated. possibly through a model
- $\text{pr}(R_i=1 \mid C_{\text{MAR}}) = \pi(C_{\text{MAR}})$
- Can specify as logistic regression
 $\text{logit}\{\pi(C_{\text{MAR}})\} = \beta^T C_{\text{MAR}}$
- Could specify any model

Modeling Choices

Step 3

- MAR assumption dictates C_{MAR} variables
- But not the form of $\text{pr}(R_i=1 \mid C_{\text{MAR}}) = \pi(C_{\text{MAR}})$
- Include transformation, spline, interactions?
lab last week needed an interaction
- Congeniality w/ analysis model not needed
interaction in one doesn't require it the other
- Validity requires $\pi(C_{\text{MAR}})$ model is correct
suggests non-parsimonious modeling

Calculate $\pi(C_{MAR})$

Step 3

- Estimable for anyone with C_{MAR} available
- Key to the unbiased analysis
- weight: $w_i = 1 / \pi_i$
 - `logit R i.hlamat age_don age death txttype year`
 - `predict pi, p`
 - `gen w = 1/p`

Sampling Weights

Step 3

- Attempts to create the full sample
- “Re-weight” Data | $R=1$ to mimic it
- Large weight $\Rightarrow$ increased importance
- When π_i small: person under-represented
leads to a large weight
- When $\pi_i = 1$, full represented, $w_i = 1$
given no additional weight in analysis

Equivalent Models

- Complete Data:

$$\begin{aligned}\text{logit}(Y) = & \beta_0 + \beta_{\text{age}} * \text{age} + \beta_{\text{age_don}} * \text{age_don} \\ & + \beta_{\text{year}} * \text{year} + \beta_{\text{txtype}} * \text{txtype} \\ & + \beta_{\text{cold_isc}} * \text{cold_isc} + \beta_{\text{hlamat}} * \text{hlamat}\end{aligned}$$

- Data with R=1

$$\begin{aligned}\text{logit}(Y) = & \beta_0 + \beta_{\text{age}} * \text{age} + \beta_{\text{age_don}} * \text{age_don} \\ & + \beta_{\text{year}} * \text{year} + \beta_{\text{txtype}} * \text{txtype} \\ & + \beta_{\text{cold_isc}} * \text{cold_isc} + \beta_{\text{hlamat}} * \text{hlamat}\end{aligned}$$

weighted by w_i

Analysis Commands in Stata

- `logit R i.hlamat age_don age death txttype year`
- `predict pi, p`
- `gen w = 1/pi`
- `logit death i.hlamat age_don age cold txttype year [pweight=w]`

Stata Results

```

Logistic regression
Log pseudolikelihood = -1717.388
Number of obs      =      7,347
Wald chi2(11)      =     131.68
Prob > chi2         =      0.0000
Pseudo R2          =      0.0436
  
```

death	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
hlamat						
1	.0962106	.2291147	0.42	0.675	-.3528461	.5452672
2	.0084009	.2333768	0.04	0.971	-.4490093	.465811
3	-.1779645	.2334702	-0.76	0.446	-.6355577	.2796286
4	-.216194	.275066	-0.79	0.432	-.7553134	.3229255
5	-.0910452	.3676077	-0.25	0.804	-.811543	.6294525
6	-.1018275	.3715002	-0.27	0.784	-.8299546	.6262995
age_don	-.0023699	.0044193	-0.54	0.592	-.0110315	.0062916
age	-.0329875	.0111418	-2.96	0.003	-.0548251	-.01115
cold_isc	.0007704	.006628	0.12	0.907	-.0122202	.0137609
txtype	.2828537	.1976179	1.43	0.152	-.1044703	.6701777
year	-.1467212	.0152131	-9.64	0.000	-.1765384	-.116904
_cons	290.0792	30.36358	9.55	0.000	230.5677	349.5908

Advice: Inverse Weights

- Use non-parsimonious model for R
- Use external information
to guide the MAR assumption/model
- Use when only a single missing variables
avoid if missing data in > 1 variable/predictor
- Examine the probability weights
large weights yield big influence
- Analysis so similar to complete data
just with weights
- Can be inefficient

Maximum Likelihood (ML)

Based on implication #1

- Maximum likelihood valid for MAR data
as long as CMAR variables incorporated
- Works especially well for missing outcome
especially in longitudinal analysis
where attrition depends on previous values
- ML often fit with the EM algorithm
EM also iterative fills in missing data
- Imputation & analysis models consistent
MI and maximum like very similar

Disadvantages of ML

- Cumbersome for missing predictors don't usual model their distribution
- Awkward if CMAR variables are not a part of desired
- Lacks flexibility of MI

3 Methods

- Same MAR assumption
- Inverse weight:
model missingness, data model flexible
- MI/ML:
model data, missingness model unspecified