

# Functional Form:

## *Modeling Continuous Predictors*

Greenland, 1995  
Altman and Royston, 2006

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20 September 2016

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## Lee Example

- Cohort of community-dwelling elders
- Predictors: age, co-morbidities, functional status: > 48 predictors
- Follow-up: mortality over 4 years
- Logistic regression model
- Age is a continuous variable

Lee, SJ et al. JAMA. 2006;295:801-808.

# Summary Statistics

summ age, detail

age at time of 1998 interview

| ----- |             |          |             |          |
|-------|-------------|----------|-------------|----------|
|       | Percentiles | Smallest |             |          |
| 1%    | 51          | 50       |             |          |
| 5%    | 53          | 50       |             |          |
| 10%   | 55          | 50       | Obs         | 11701    |
| 25%   | 59          | 50       | Sum of Wgt. | 11701    |
| 50%   | 66          |          | Mean        | 67.22169 |
|       |             | Largest  | Std. Dev.   | 10.10853 |
| 75%   | 75          | 101      |             |          |
| 90%   | 82          | 101      | Variance    | 102.1825 |
| 95%   | 85          | 102      | Skewness    | .4472938 |
| 99%   | 92          | 103      | Kurtosis    | 2.458463 |

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## Distribution of Ages

- Looking for outliers
- Look to see where most values are  
*90% of ages 53 and 85*
- *DONT CARE IF THEIR NORMALLY DIST  
doesn't matter*
- *It doesn't.*

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# Modeling Challenge

## Levels of Complexity

- age less than <65
- age 50-60, 60-70, 70-80, 80-90, 90+
- age as a “linear term”
- age + age<sup>2</sup>
- other transformation (e.g. log)
- smoothing

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## Age Cut at 65

Logistic regression

Log likelihood = -3910.362

|               |   |        |
|---------------|---|--------|
| Number of obs | = | 11701  |
| LR chi2(1)    | = | 592.71 |
| Prob > chi2   | = | 0.0000 |
| Pseudo R2     | = | 0.0704 |

| died   | Odds Ratio | Std. Err. | z     | P> z  | [95% Conf. Interval] |          |
|--------|------------|-----------|-------|-------|----------------------|----------|
| over65 | 4.829268   | .3501685  | 21.72 | 0.000 | 4.189489             | 5.566749 |

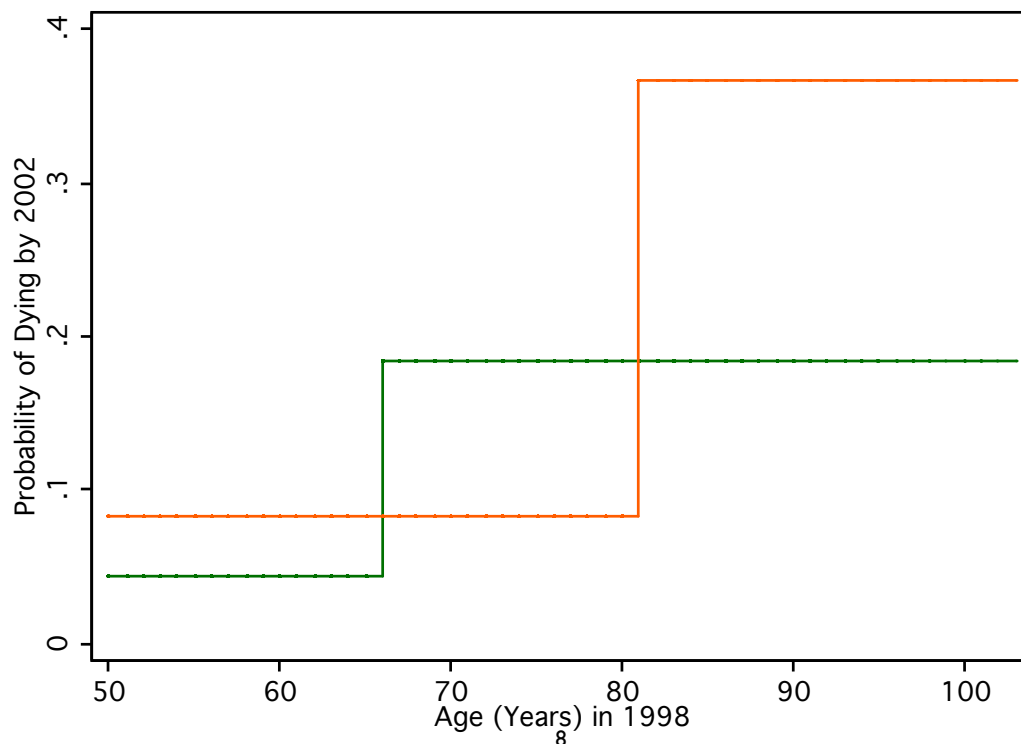
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# Odds Ratios

| Age 1 | Age 2 | Odds Ratio |
|-------|-------|------------|
| 60    | 64    | 1          |
| 64    | 66    | 4.8        |
| 71    | 79    | 1          |
| 79    | 81    | 1          |

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## Two Fits



# How to Choose Cutpoint

- Strongly discourage making binary variable
- Lose a lot of information
- Cutpoint can be important choice
  - data: no unambiguous choice
  - avoid choosing “most significant” cutpoint

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## More Levels

- Cut age at quartiles: 25, 50, 75  
*ages 59, 66, 75*
- Use as a 4-level ordinal categorical variable
- Why use trend test?

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# Age by Quartile

```
. xi: logistic died i.age4
i.age4          _Iage4_1-4      (naturally coded; _Iage4_1 omitted)

Logistic regression                                Number of obs   =      11701
                                                    LR chi2(3)       =      949.71
                                                    Prob > chi2      =      0.0000
Log likelihood = -3731.8609                        Pseudo R2       =      0.1129
```

| died     | Odds Ratio | Std. Err. | z     | P> z  | [95% Conf. Interval] |          |
|----------|------------|-----------|-------|-------|----------------------|----------|
| _Iage4_2 | 2.299158   | .2939617  | 6.51  | 0.000 | 1.789523             | 2.95393  |
| _Iage4_3 | 3.842798   | .4605231  | 11.23 | 0.000 | 3.038359             | 4.860222 |
| _Iage4_4 | 12.90059   | 1.454704  | 22.68 | 0.000 | 10.3425              | 16.09137 |

Learn anything about shape?

||

## Interpretation

- OR is 12.9 comparing 4th quartile to first often need that 1 number
- Common summary for continuous *especially in NEJM*
- Quartile can be very specific to your data

# Odds Ratios

| Age 1 | Age 2 | Odds Ratio |
|-------|-------|------------|
| 60    | 64    | 1          |
| 64    | 66    | 1.68       |
| 71    | 79    | 3.34       |
| 79    | 81    | 1          |

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## Issues

### Binary and Categorical

- Con:
  - potentially biased
  - subtle choices
  - residual confounding
- Pro
  - very simple to understand
  - easy to “act” on cutpoints

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# Continuous Predictor

```
. logistic died age
```

Logistic regression

Log likelihood = -3632.9744

Number of obs = 11701

LR chi2(1) = 1147.48

Prob > chi2 = 0.0000

Pseudo R2 = 0.1364

| died | Odds Ratio | Std. Err. | z     | P> z  | [95% Conf. Interval] |          |
|------|------------|-----------|-------|-------|----------------------|----------|
| age  | 1.103388   | .0034619  | 31.36 | 0.000 | 1.096623             | 1.110194 |

Interpretation?

Scale is huge issue here

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# Changing Scale

```
lincom 5*age  
( 1) 5 age = 0
```

| died | Odds Ratio | Std. Err. | z     | P> z  | [95% Conf. Interval] |          |
|------|------------|-----------|-------|-------|----------------------|----------|
| (1)  | 1.635463   | .0256563  | 31.36 | 0.000 | 1.585942             | 1.686529 |

```
lincom 10*age  
( 1) 10 age = 0
```

| died | Odds Ratio | Std. Err. | z     | P> z  | [95% Conf. Interval] |         |
|------|------------|-----------|-------|-------|----------------------|---------|
| (1)  | 2.674738   | .08392    | 31.36 | 0.000 | 2.515213             | 2.84438 |

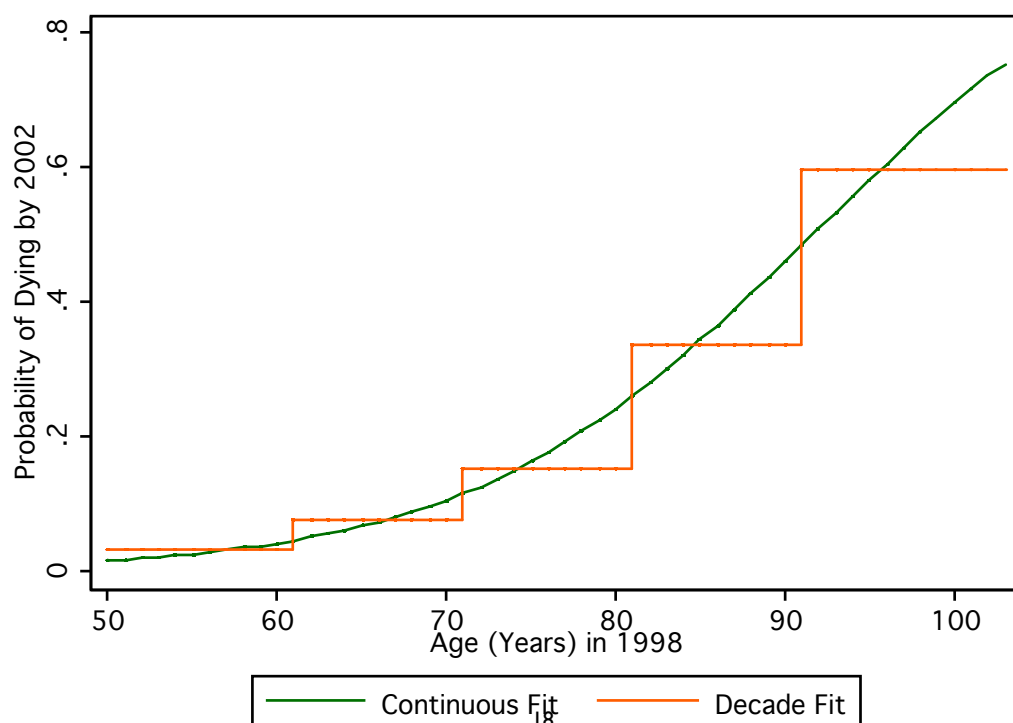
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# Induced Odds Ratios

| Age 1 | Age 2 | Odds Ratio |
|-------|-------|------------|
| 60    | 64    | 1.48       |
| 64    | 66    | 1.22       |
| 71    | 79    | 2.2        |
| 79    | 81    | 1.22       |

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## Linear v. Decade Fit



# Linear Fit

- Can give good easy predictions
- Requires only single df: parsimonious
- Avoids overfitting!!  
not big concern in this data
- Bias/variance tradeoff  
may give better predictor for very old
- Single parameter output  
every 10 year increases odds 2.67

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## Does the Model Fit

- Maybe the very simple model is missing a pattern
- Multiple categories don't have to give monotone risk
- One way to check: squared term  
*squared terms have problems*

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# Quadratic Fit

```
. logistic died age age2
```

Logistic regression

Number of obs = 11701

LR chi2(2) = 1151.82

Prob > chi2 = 0.0000

Log likelihood = -3630.8075

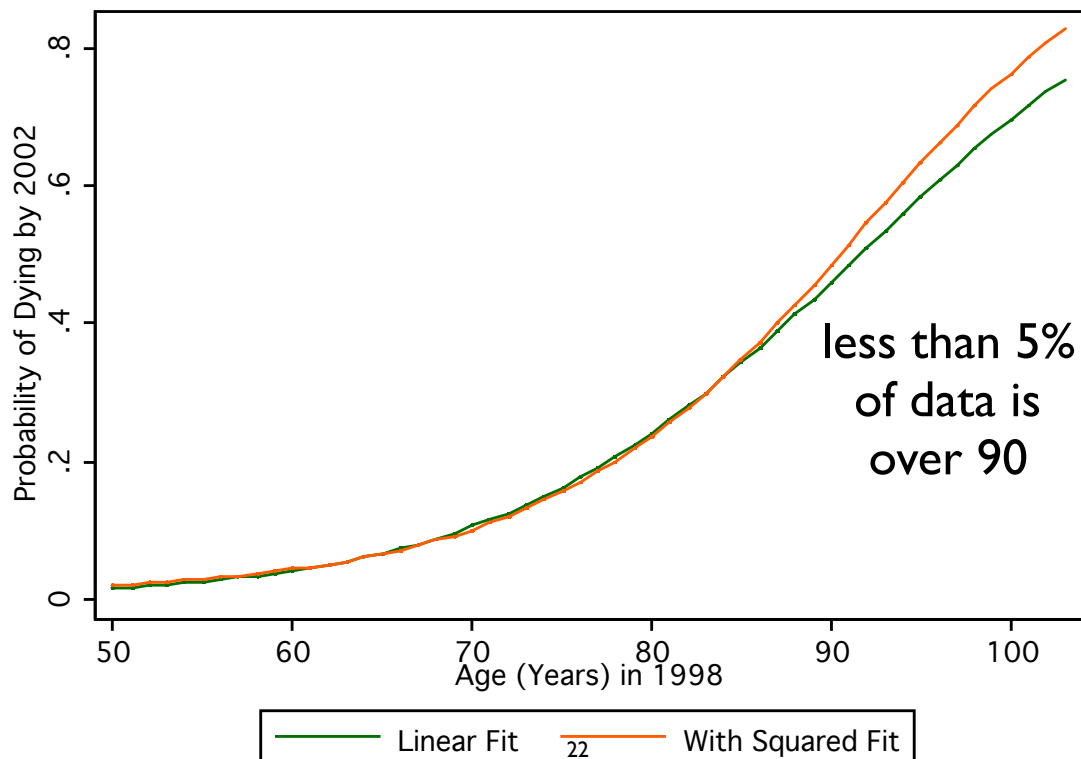
Pseudo R2 = 0.1369

| died | Odds Ratio | Std. Err. | z    | P> z  | [95% Conf. Interval] |          |
|------|------------|-----------|------|-------|----------------------|----------|
| age  | 1.016908   | .0397754  | 0.43 | 0.668 | .9418631             | 1.097932 |
| age2 | 1.000557   | .0002666  | 2.09 | 0.037 | 1.000034             | 1.001079 |

Squared term “adds” to the model  $p < 0.05$

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# Difference in Fits



# Linear v. Quadratic Fit

| Age 1 | Age 2 | OR Linear | OR Quad |
|-------|-------|-----------|---------|
| 60    | 64    | 1.48      | 1.41    |
| 64    | 66    | 1.22      | 1.2     |
| 71    | 79    | 2.2       | 2.23    |
| 79    | 81    | 1.22      | 1.24    |

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## Squared Terms

- Very prone to outliers  
*outliers become magnified*  
*very non-robust models*  
*can affect fit far from outlying values*
- Statistical v. practical significance in model improvement

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# Log transformation

- Does not add degrees of freedom per se
- But can handle “non-linear” patterns
- Produces a useful interpretation
- Interpretation in terms of % increase in predictor
- Particularly helpful when measurement is unfamiliar

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## Results of Log Fit

```
. logistic died lage
```

Logistic regression

```
Number of obs   =    11701
LR chi2(1)      =    1132.03
Prob > chi2     =    0.0000
Pseudo R2      =    0.1346
```

Log likelihood = -3640.7019

| died | Odds Ratio | Std. Err. | z     | P> z  | [95% Conf. Interval] |         |
|------|------------|-----------|-------|-------|----------------------|---------|
| lage | 1166.966   | 270.697   | 30.44 | 0.000 | 740.6419             | 1838.69 |

Interpretation problematic

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# Results of Log Fit

```
. lincom 0.095*lage
( 1)  .095 lage = 0
```

| died | Odds Ratio | Std. Err. | z     | P> z  | [95% Conf. Interval] |          |
|------|------------|-----------|-------|-------|----------------------|----------|
| (1)  | 1.956008   | .0431042  | 30.44 | 0.000 | 1.873324             | 2.042341 |

10% increase in age  
 same as multiplying age by 1.1  
 same as  $\log(1.1) = 0.095$  increase in log age

10% increase in age => HR 1.96

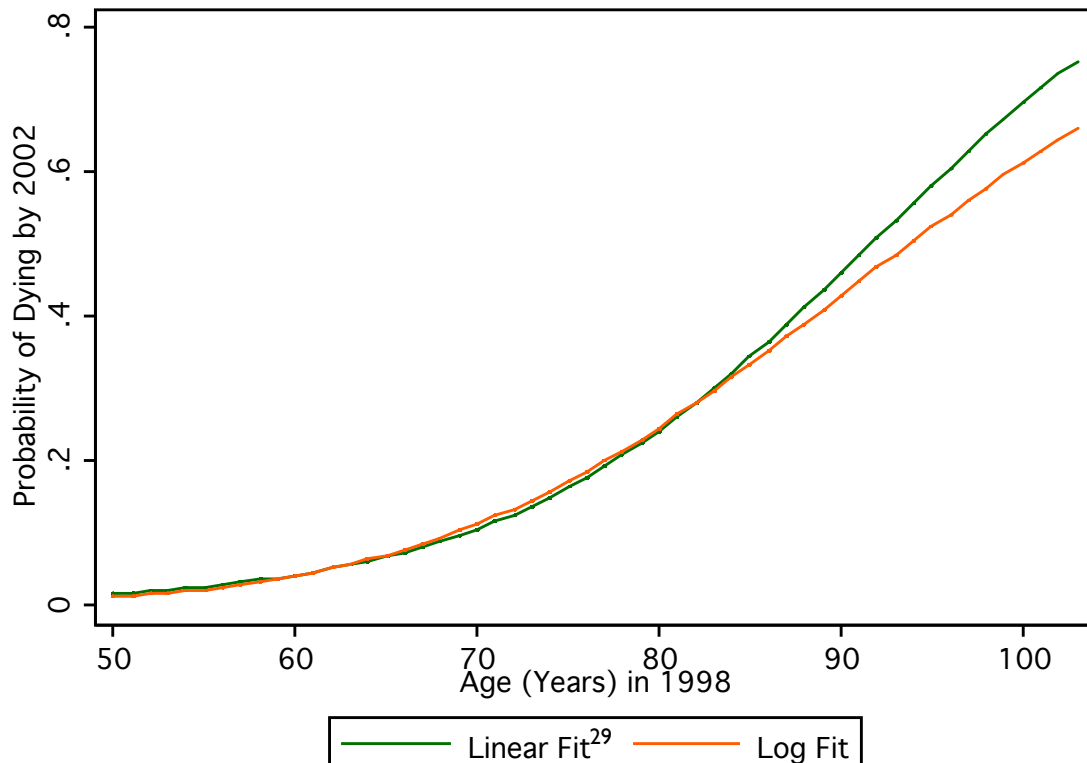
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## Log v. Linear Fit

| Age 1 | Age 2 | OR Log | OR Linear |
|-------|-------|--------|-----------|
| 60    | 64    | 1.58   | 1.48      |
| 64    | 66    | 1.24   | 1.22      |
| 71    | 79    | 2.12   | 2.2       |
| 79    | 81    | 1.19   | 1.22      |

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# Linear v. Log Fit



## Consider log transform

- Log likelihood reveals no improvement in fit
- In addition, interpretation not helpful
- Sometimes hear “twice your age”
- Generally talk about being 1,5, etc years older
- Additive scale more natural

# Splines

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# Splines

- A tool for producing smooth functions
- Very amenable to regression models
- Can help in describing complex shape *while adjusting for other variables*
- Similar across regression models

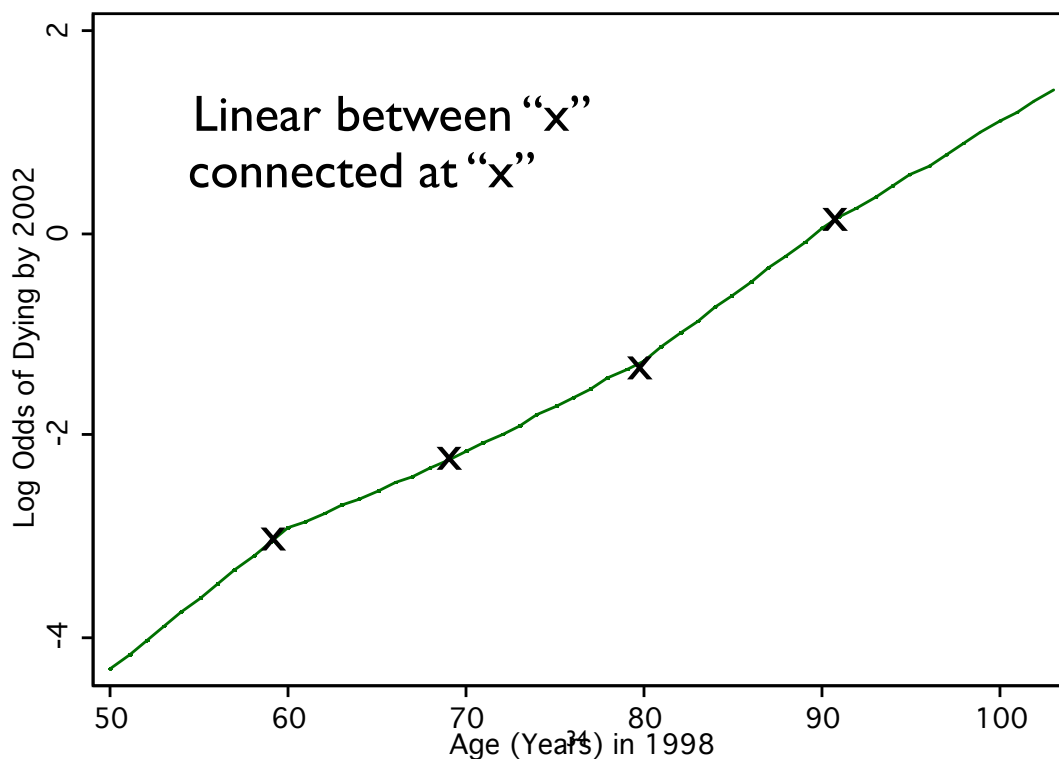
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# Heuristic

- Approximate complex function with
- Multiple simple functions
- Simplest example: linear spline
- Suppose use simple lines to approximate effect of age

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## Linear Spline



# Knots

- “x” are called knots  
control how much shape can take
- Where to put knots?
- How many to put down?
- First, not a big effect
- Later can have big effect

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## # of Knots

- Controls how eccentric curve can be
- $DF = \# \text{ Knots} - 1$
- Too many knots: overfitting  
noisy curve
- Too few: bias
- Bias/variance trade off
- 3-5 knots is plenty

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# Advantage Over Cutpoint

- Get much more information for df
- Placement of points not as critical
- Extremely flexible methods

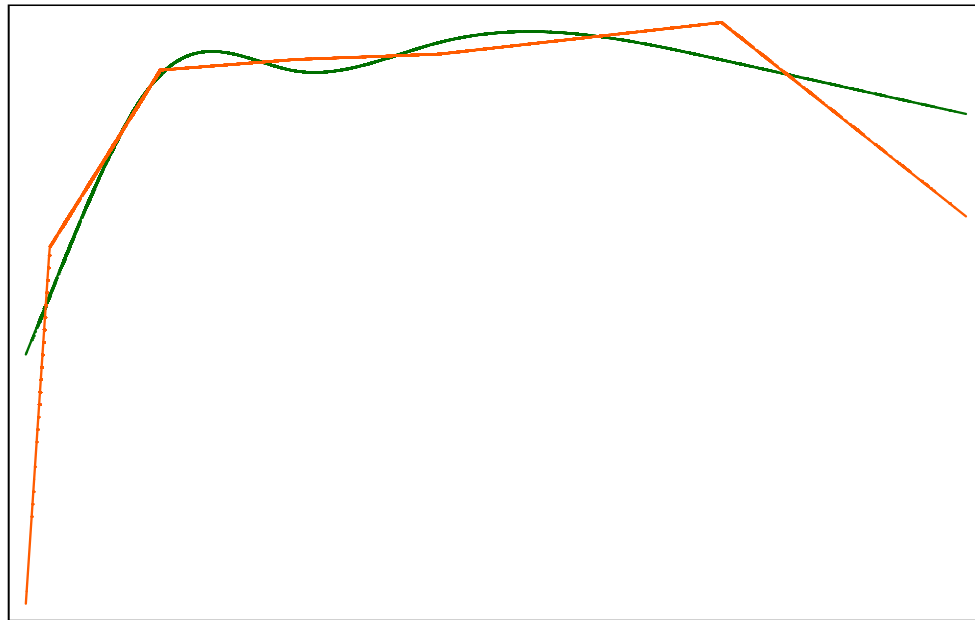
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## Cubic Splines

- Type of spline functions
- Fits cubic polynomial between knots  
*more flexible than a line*
- Local function: not susceptible to outliers
- Often need to constrain tails
- Produces nice, smooth looking functions

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# Cubic v. Linear Spline



— Cubic Spline — Linear Spline

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## Stata

```
. mkspline cube_age = age, cubic nknots(3) displayknots
```

|     | knot1 | knot2 | knot3 |
|-----|-------|-------|-------|
| age | 55    | 66    | 82    |

```
. summ cube_age*
```

| Variable  | Obs   | Mean     | Std. Dev. | Min | Max      |
|-----------|-------|----------|-----------|-----|----------|
| cube_age1 | 11701 | 67.22169 | 10.10853  | 50  | 103      |
| cube_age2 | 11701 | 5.493749 | 7.357163  | 0   | 43.18518 |

# Model Fit

```
. logistic died cube_age*
```

Logistic regression

Number of obs = 11701

LR chi2(2) = 1149.94

Prob > chi2 = 0.0000

Log likelihood = -3631.7479

Pseudo R2 = 0.1367

| died      | Odds Ratio | Std. Err. | z    | P> z  | [95% Conf. Interval] |          |
|-----------|------------|-----------|------|-------|----------------------|----------|
| cube_age1 | 1.084836   | .0120655  | 7.32 | 0.000 | 1.061444             | 1.108743 |
| cube_age2 | 1.019988   | .0128014  | 1.58 | 0.115 | .9952037             | 1.045389 |

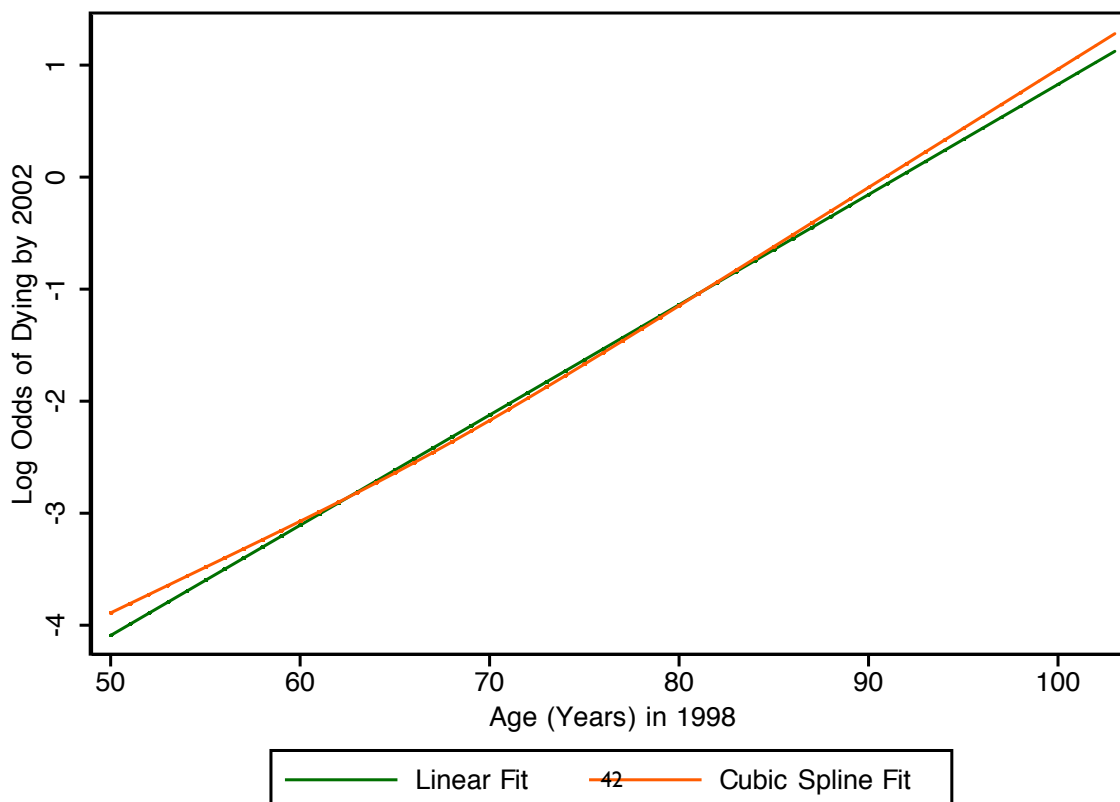
coefficient not interpretable

extract prediction

```
predict out_cubic
```

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# Spline Fit



# How to test splines

*flexible test for non-linear effects*

- Fit spline model  
`logistic died cube_age1 cube_age2`
- Store results  
`est store splineresults`
- Fit model with linear effect  
`logistic died age`
- Apply likelihood ratio test  
`lrtest splineresults`
- LR test works because spline model always nested

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# How to test splines

*simpler alternative for cubic splines*

- Fit spline model  
`logistic died cube_age1 cube_age2`
- Note, “first” term is the linear effect  
*first term:* `cube_age1`
- Subsequent terms allow for non-linearity  
*here cubic there is only* `cube_age2`
- Test that subsequent terms 0  
`test cube_age2 = 0`

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# Two Tests

## Likelihood Ratio Test

```
. lrtest splineresults
```

|                                         |               |        |
|-----------------------------------------|---------------|--------|
| Likelihood-ratio test                   | LR chi2(1) =  | 2.45   |
| (Assumption: . nested in splineresults) | Prob > chi2 = | 0.1173 |

## Wald Test

```
. test cube_age2=0
```

```
( 1) [died]cube_age2 = 0
```

|               |        |
|---------------|--------|
| chi2( 1) =    | 2.49   |
| Prob > chi2 = | 0.1148 |

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# Adjusted Splines

- Predicted outcome (prob, mean) by cts predictor
- All other variables “held constant”
- Race adjusted mortality: requires we specify racial composition of ref. population
- Requires that we margin race out of predictions
- According to mix in reference population

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# Race in Sample

| race/ethnicity  | Freq.  | Percent | Cum.   |
|-----------------|--------|---------|--------|
| white, non-hisp | 9,467  | 80.99   | 80.99  |
| black, non-hisp | 1,225  | 10.48   | 91.47  |
| other, non-hisp | 225    | 1.92    | 93.40  |
| hisp            | 772    | 6.60    | 100.00 |
| Total           | 11,689 | 100.00  |        |

I set reference population to resemble this: 81% white, 10% black, 2% other, 6% latino

Choice is *arbitrary* -- could have chose race mix for those under 65. Or the pop of California

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# Adjusted Model

```
. logistic died cube_age* i.raceeth
```

```
Logistic regression               Number of obs   =      11689
                                LR chi2(5)        =      1165.06
                                Prob > chi2        =       0.0000
Log likelihood = -3618.6447       Pseudo R2      =       0.1387
```

| died      | Odds Ratio | Std. Err. | z      | P> z  | [95% Conf. Interval] |          |
|-----------|------------|-----------|--------|-------|----------------------|----------|
| cube_age1 | 1.086678   | .0121674  | 7.42   | 0.000 | 1.06309              | 1.110789 |
| cube_age2 | 1.018119   | .0128472  | 1.42   | 0.155 | .9932481             | 1.043613 |
| raceeth   |            |           |        |       |                      |          |
| 2         | 1.416758   | .1341638  | 3.68   | 0.000 | 1.176762             | 1.7057   |
| 3         | .9241177   | .22599    | -0.32  | 0.747 | .5722275             | 1.492402 |
| 4         | .9216928   | .1313255  | -0.57  | 0.567 | .6971151             | 1.218619 |
| _cons     | .0003029   | .0002135  | -11.49 | 0.000 | .0000761             | .0012061 |

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# Key Code

```
gen adjlogodds = _b[_cons] + _b[cube_age1]*cube_age1 +  
_b[cube_age2]*cube_age2 + _b[1.raceeth]*0.80 _b[2.raceeth]*0.10 +  
_b[3.raceeth]*0.02 + _b[3.raceeth]*0.07
```

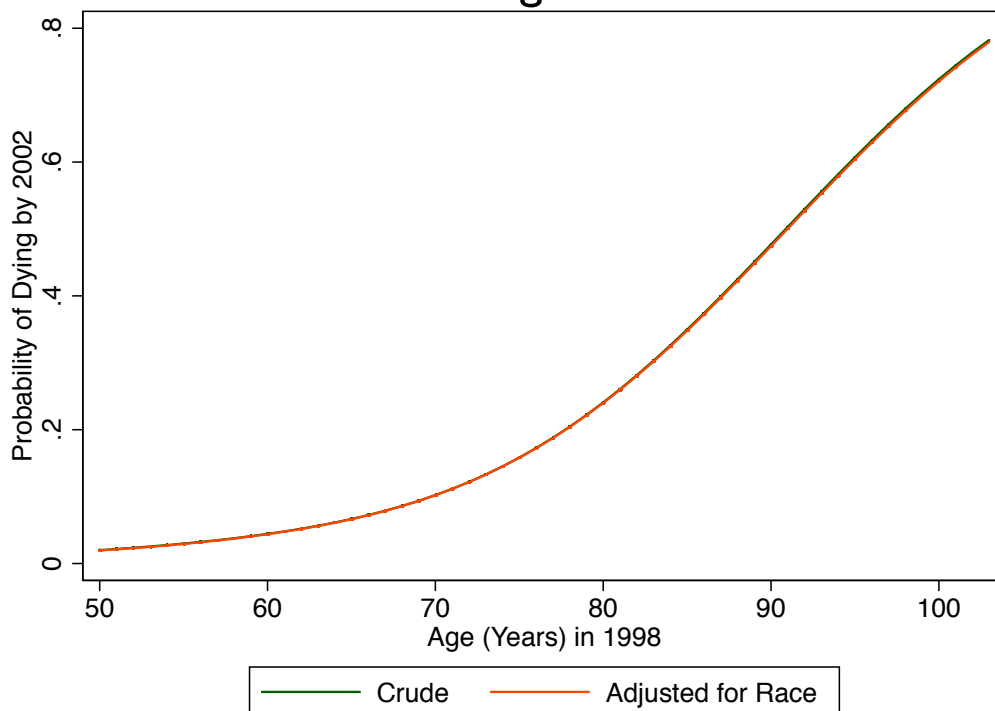
Calculates Logs Odds for the Reference Population

```
gen adjprob = exp( adjlogodds ) / (1+ exp( adjlogodds ))
```

Converts to a Predicted Probability

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## Crude v. Adjusted Indistinguishable



# Cutpoint Approach

- Binary cutpoint approach always inferior.  
*Can always do better*
- Multiple cutpoints: better
- Dilemma: flexibility v. degrees of freedom
- Discontinuous: difficult for statisticians
- Groups: appeals to clinicians
- Choice of cutpoint: difficult and important
- Residual confounding is an issue

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# Transformations:

## Polynomial

- Two df but only moderate flexibility
- Very susceptible to outliers
- Outlier can affect curve far away
- Mathematically complex: need to graph
- Statistical v. practical significance

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# Transformations:

## Log Transformation

- Pulls in outliers
- Can produce useful interpretation  
*especially for exotic assays*
- Mathematically complex
- Don't need to graph
- Can't handle negative values
- No easy test for linear v. log

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## Linear Fit

- Parsimonious
- No worry about overfitting  
*optimizing bias/variance tradeoff*
- Often “good enough”  
*especially to get rid of confounding*
- Need to be careful about the scale
- Usually produces good interpretation  
*in the eye of the beholder*

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# Spline: Pros

- Flexible method
- Knot placement not critical
- Can be fit in any regression model
- Allows for check on linear fit
- Potential rich information  
*especially for degrees of freedom spent*

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# Spline: Cons

- Knot number is important issue
- Too many knots can lead to overfitting  
*could use cross-validation to choose*
- Can only be conveyed in a plot  
*difficult when adjusting other variables*
- Useful when predictor is of HIGH interest

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# Balance

- overfitting (-)
- residual confounding (-)
- flexibility (+)
- interpretation (+)

Purpose of regression very important