

Biostat 200

Lecture 11

Today

- Linear regression: Categorical explanatory variables
- Multiple regression
- Logistic regression
- Wrap up
- Please do evaluations on moodle

From last time

- Plotting residuals in Stata
 - regress fev ht
 - rvfplot ** gives you Residuals vs. Fitted
 - rvpplot ht ** gives you Residuals vs. Predictor
(indep var)

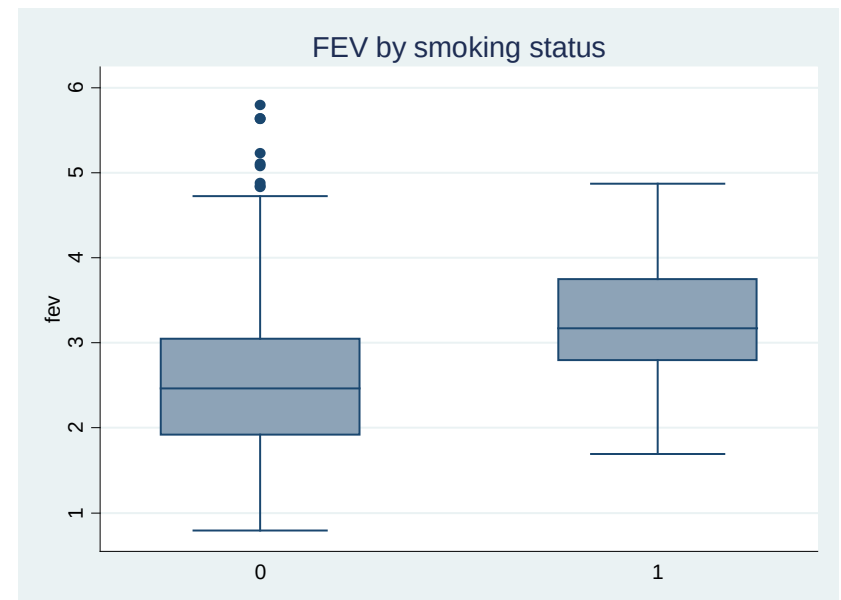
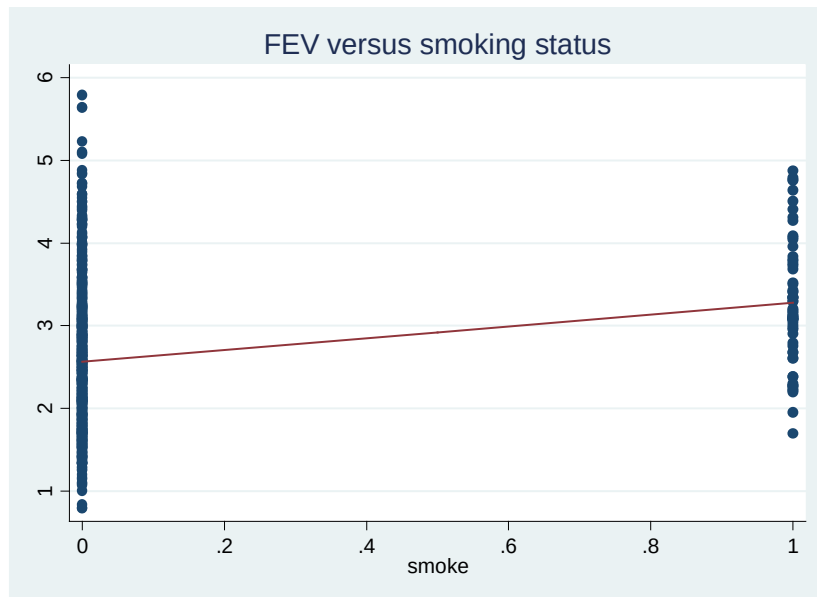
Linear regression with categorical explanatory variable

```
. regress fev smoke
```

Source	SS	df	MS	Number of obs = 654		
Model	29.569683	1	29.569683	F(1, 652) = 41.79		
Residual	461.35015	652	.707592255	Prob > F = 0.0000		
Total	490.919833	653	.751791475	R-squared = 0.0602		
				Adj R-squared = 0.0588		
				Root MSE = .84119		

fev	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
smoke	.7107189	.1099426	6.46	0.000	.4948346	.9266033
_cons	2.566143	.0346604	74.04	0.000	2.498083	2.634202

Regression with categorical predictors



- There is a lot of variability in each group here
- Smoking yes/no may be a crude measure

Multiple regression

- Additional explanatory variables might add to our understanding of a dependent variable
- We can posit the population equation

$$\mu_{y/x_1, x_2, \dots, x_q} = \alpha + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_q x_q$$

- α is the mean of y when all the explanatory variables are 0
- β_i is the change in the mean value of y that corresponds to a 1 unit change in x_i when all the other explanatory variables are held constant

- Because there is natural variation in the response variable, the model we fit is

$$y = \alpha + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_q x_q + \varepsilon$$

- Assumptions
 - $x_1, x_2, \dots, x_q$ are measured without error
 - The distribution of y is normal with mean $\mu_{y/x_1, x_2, \dots, x_q}$ and standard deviation $\sigma_{y/x_1, x_2, \dots, x_q}$
 - The population regression model holds
 - For any set of values of the explanatory variables, $x_1, x_2, \dots, x_q$, $\sigma_{y/x_1, x_2, \dots, x_q}$ is constant – homoscedasticity
 - The outcomes are independent

You can fit continuous and/or categorical variables

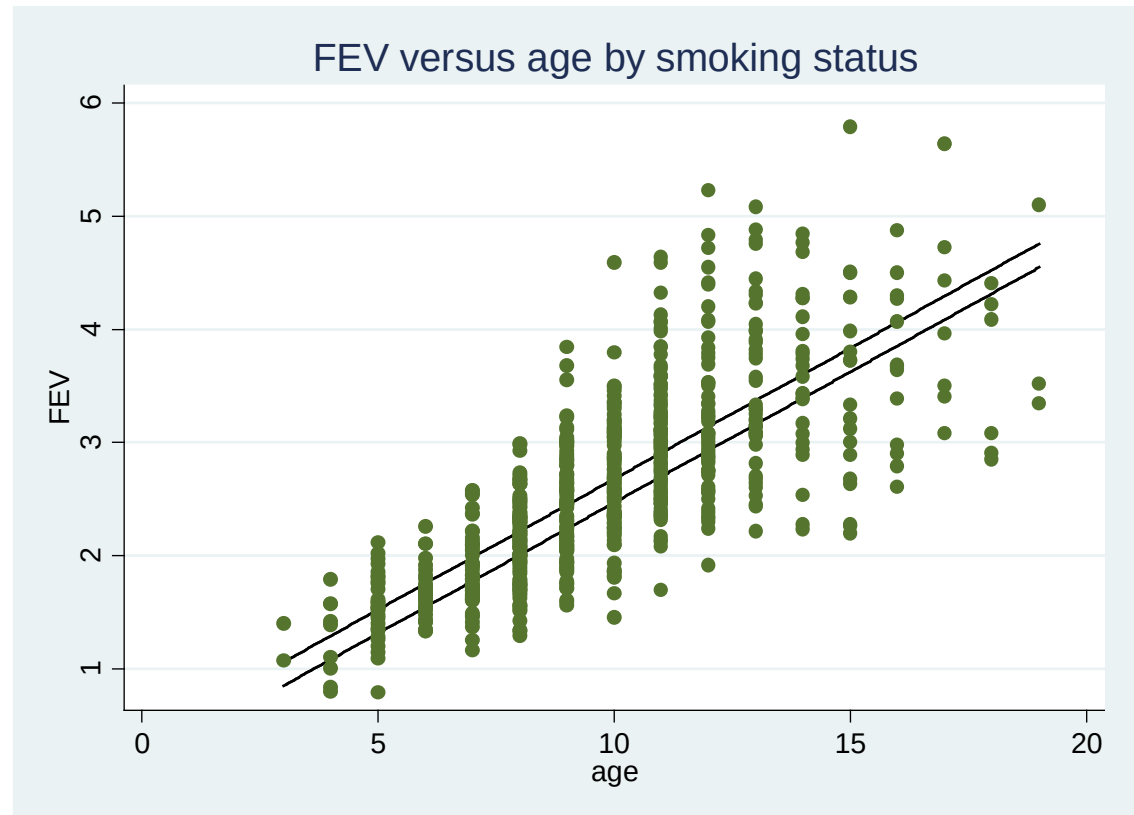
```
. regress fev age smoke
```

Source	SS	df	MS	Number of obs	=	654
Model	283.058247	2	141.529123	F(2, 651)	=	443.25
Residual	207.861587	651	.319295832	Prob > F	=	0.0000
Total	490.919833	653	.751791475	R-squared	=	0.5766
				Adj R-squared	=	0.5753
				Root MSE	=	.56506

fev	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	.2306046	.0081844	28.18	0.000	.2145336	.2466755
smoke	-.2089949	.0807453	-2.59	0.010	-.3675476	-.0504421
_cons	.3673731	.0814357	4.51	0.000	.2074647	.5272814

- The model is $\hat{fev} = \hat{\alpha} + \beta_1 \text{ age} + \beta_2 X_{\text{smoke}}$
- So for non-smokers, we have $\hat{fev} = \hat{\alpha} + \beta_1 \text{ age}$
- For smokers, $\hat{fev} = \hat{\alpha} + \beta_1 \text{ age} + \beta_2$
 - So β_2 is the mean difference in FEV for smokers versus non-smokers at each age

- When you have one continuous variable and one dichotomous variable, you can think of fitting two lines that only differ in y intercept by the coefficient of the dichotomous variable
- E.g. $\beta_2 = -.209$



Logistic regression

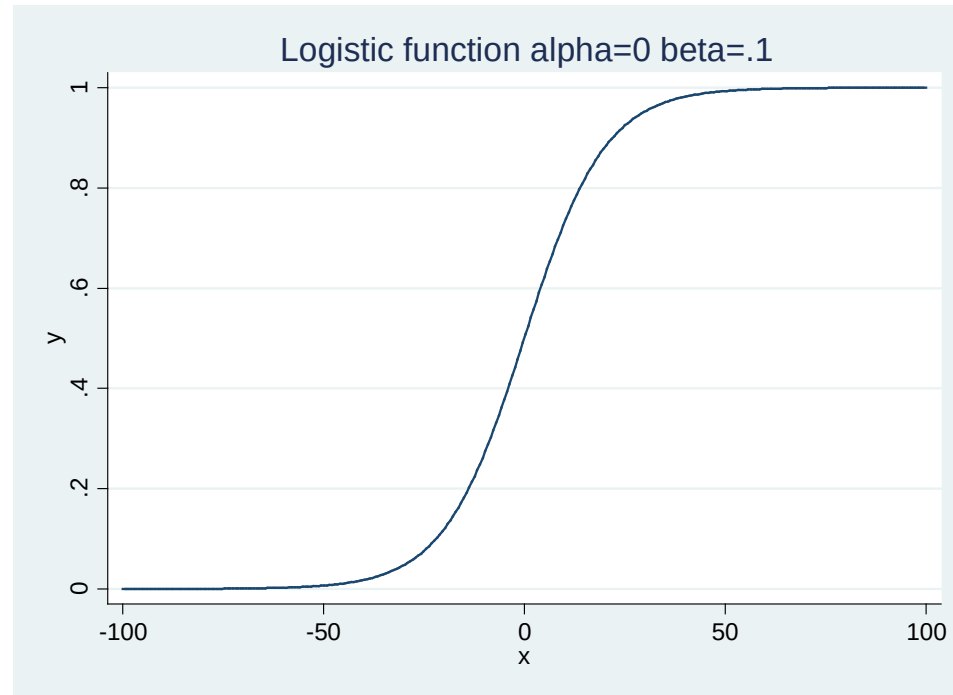
- For linear regression
 - The predicted values can be from $-\infty$ to $+\infty$
 - Y values are assumed to follow a normal distribution
- For many situations we want to model a dichotomous outcome, such as disease/no disease and cannot use linear regression
- We want to model p , the probability of disease

Logistic regression

- $p = P(Y=1)$ i.e. $P(\text{outcome}=\text{disease})$
- A model of the form $p = \alpha + \beta x$ would be able to take on negative values or values more than 1
- $p = e^{\alpha + \beta x}$ is an improvement because it cannot be negative, but it still could be greater than 1

Logistic regression

- How about the function?
$$p = \frac{e^{\alpha + \beta x}}{1 + e^{\alpha + \beta x}}$$



- This function =.5 when $\alpha + \beta x = 0$
- The function models the probability slowly increasing over the values of X, until there is a steep rise, and another leveling off

Logistic regression

- The logistic function
$$p = \frac{e^{\alpha + \beta x}}{1 + e^{\alpha + \beta x}}$$

- Now check this out:

$$1 - p = 1 - \frac{e^{\alpha + \beta x}}{1 + e^{\alpha + \beta x}} = \frac{1 + e^{\alpha + \beta x} - e^{\alpha + \beta x}}{1 + e^{\alpha + \beta x}} = \frac{1}{1 + e^{\alpha + \beta x}}$$

- So odds of disease/success are

$$\frac{p}{1 - p} = \frac{\frac{e^{\alpha + \beta x}}{1 + e^{\alpha + \beta x}}}{\frac{1}{1 + e^{\alpha + \beta x}}} = e^{\alpha + \beta x}$$

- And therefore $\ln(p/(1-p)) = \alpha + \beta x$

Logistic regression

- So instead of assuming that the relationship between X and p is linear , we are assuming that the relationship between $\ln(p/(1-p))$ and X is linear.
- $\ln(p/(1-p))$ is called the logit function

Logistic regression

- This transformation of the outcome p is called the logit link
- It is part of a family of models called generalized linear models
- In generalized linear models, some function of the outcome variable is assumed to be linear

e.g. $f(y) = \alpha + bx$ or $f(p) = \alpha + bx$

e.g. $\text{logit}(p) = \alpha + bx$

Logistic regression

- Assumptions
 - Y_i follows a bernoulli distribution (0,1)
 - The mean of Y at every X , $P(Y=1 | x)$ is given by the logistic function
 - The values of the outcome are independent

Logistic regression

- Estimation
 - The parameters are estimated via a procedure called *maximum likelihood*
 - There is not a closed form solution of the maximization (i.e. we cannot write down a solution) – it is an iterative procedure

Notation

- $p_0 = P(Y=1 | X=0)$
is the probability that $y=1$ when $x=0$
- $p_1 = P(Y=1 | X=1)$
is the probability that $y=1$ when $x=1$
- $p = P(Y=1)$
is the probability that $y=1$ for any x

Interpretation of coefficients

- One dichotomous explanatory variable
- For $x=0$

$$\begin{aligned} & \ln(P(Y=1 | X=0)/(1- P(Y=1 | X=0))) \\ &= \ln(p_0/(1-p_0)) \\ &= \alpha + \beta * 0 = \alpha \end{aligned}$$

- For $x=1$

$$\begin{aligned} & \ln(P(Y=1 | X=1)/(1- P(Y=1 | X=1))) \\ &= \ln(p_1/(1-p_1)) \\ &= \alpha + \beta * 1 = \alpha + \beta \end{aligned}$$

Interpretation of coefficients

$$\ln \left(\frac{\frac{p_1}{1-p_1}}{\frac{p_0}{1-p_0}} \right) = \ln(OR)$$

- $\ln[(p_1/(1-p_1))/(p_0/(1-p_0))]$
= $(\ln(p_1/(1-p_1)) - \ln(p_0/(1-p_0))) =$
= $(\alpha + \beta) - \alpha = \beta$

→ $OR = e^\beta$

*Remember that $\ln(a/b) = \ln(a) - \ln(b)$

Correlates of having had a cold in the prior month

```
. logistic coldany i.rested_mostly
```

Logistic regression

Number of obs = 452

LR chi2(1) = 2.43

Prob > chi2 = 0.1188

Log likelihood = -272.65565

Pseudo R2 = 0.0044

```
-----+-----
      coldany | Odds Ratio   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
1.rested_m~y |   .6999519   .162216   -1.54   0.124   .4444262   1.102394
-----+-----
```

Correlates of having had a cold in the prior month

logistic depvar indepvar

```
. logistic coldany i.rested_mostly
```

Logistic regression

Number of obs = 452
LR chi2(1) = 2.43
Prob > chi2 = 0.1188
Pseudo R2 = 0.0044

Log likelihood = -272.65565

coldany	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]
1.rested_m~y	.6999519	.162216	-1.54	0.124	.4444262 1.102394

$=e^{\beta}$

This is used in comparing nested models of the same data set

Test of whether the model fit differs from the model with only α and no covariates

The Pseudo R^2 is an attempt to be analogous with linear regression but it is not

Stata notes

- logistic depvar indepvars
 - gives you odds ratios
- logistic depvar indepvars, coef
 - gives you coefficients (alphas and betas)
- logit depvar indepvars
 - gives you coefficients
- logit depvar indepvars, or
 - gives you odds ratios

R-square statistic in logistic regression

“In OLS regression, the R-square statistic indicates the proportion of the variability in the dependent variable that is accounted for by the model (i.e., all of the independent variables in the model). Unfortunately, creating a statistic to provide the same information for a logistic regression model has proved to be very difficult. Many people have tried, but no approach has been widely accepted by researchers or statisticians. The output from the **logit** and **logistic** commands give a statistic called "pseudo-R-square", and the emphasis is on the term "pseudo". This statistic should be used only to give the most general idea as to the proportion of variance that is being accounted for. ... There is little agreement regarding an R-square statistic in logistic regression, and that different approaches lead to very different conclusions. If you use an R-square statistic at all, use it with great care. “

<http://www.ats.ucla.edu/stat/stata/webbooks/logistic/chapter1/statalog1.htm>

Correlates of having had a cold in the prior month

```
. tabodds coldany rested_mostly, or
```

```
-----  
rested_mos~y | Odds Ratio      chi2      P>chi2      [95% Conf. Interval]  
-----+-----  
          No |      1.000000      .      .      .      .  
          Yes |      0.699952      2.38      0.1231      0.443668      1.104276  
-----  
Test of homogeneity (equal odds): chi2(1) =      2.38  
                                   Pr>chi2 =      0.1231  
  
Score test for trend of odds:      chi2(1) =      2.38  
                                   Pr>chi2 =      0.1231
```

For one independent variable that is categorical, the odds ratio estimate is the same for logistic regression as for tabular methods.

Continuous explanatory variable

```
.  
. . logistic coldany age
```

Logistic regression

Number of obs = 451
LR chi2(1) = 9.49
Prob > chi2 = 0.0021
Pseudo R2 = 0.0174

Log likelihood = -268.77615

coldany	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
-----+						
age	.9739414	.0088552	-2.90	0.004	.9567392	.9914529

- Interpretation of the coefficients: The odds ratio is for a one unit change in the predictor
- For this example the 0.974 is the odds ratio for a year difference in age

Continuous explanatory variable

- When the explanatory variable is a continuous variable, there is no equivalent tabular method without dividing the variable into groups (e.g. age groups)
- You need to use logistic regression when you want to model the relationship between a dichotomous outcome and a continuous explanatory variable

Continuous explanatory variable

- To find the OR for a 10-year change in age

.

```
. logistic coldany age, coef
```

Logistic regression

Number of obs = 451

LR chi2(1) = 9.49

Prob > chi2 = 0.0021

Log likelihood = -268.77615

Pseudo R2 = 0.0174

coldany	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
-----+-----						
age	-.0264041	.0090922	-2.90	0.004	-.0442244	-.0085838
_cons	.0902805	.338474	0.27	0.790	-.5731163	.7536773

OR for a 10-year change in age = $\exp(10 \times -.0264) = 0.768$

- To find the OR for a 10-year change in age

```
. gen age_10=age/10
(2 missing values generated)
```

```
. logistic coldany age_10
```

Logistic regression

```
Number of obs   =      451
LR chi2(1)      =       9.49
Prob > chi2     =     0.0021
Pseudo R2      =     0.0174
```

Log likelihood = -268.77615

```
-----+-----
      coldany | Odds Ratio   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
      age_10 |   .7679418   .0698225    -2.90   0.004     .6425926     .9177426
-----+-----
```



Checking logit-linearity

```
. tab agecat
```

RECODE of age	Freq.	Percent	Cum.
18-30	161	35.54	35.54
30-45	202	44.59	80.13
45-60	42	9.27	89.40
>60	48	10.60	100.00
Total	453	100.00	

```
. logistic coldany i.agecat
```

Logistic regression	Number of obs	=	451
	LR chi2(3)	=	8.45
	Prob > chi2	=	0.0375
Log likelihood = -269.29729	Pseudo R2	=	0.0154

coldany	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]
agecat					
2	1.037197	.2348894	0.16	0.872	.6654165 1.616699
3	.6011029	.2486731	-1.23	0.219	.2671845 1.352342
4	.3648981	.1615593	-2.28	0.023	.1532147 .8690459

Logistic regression with >1 explanatory variable (multiple logistic regression)

- In multiple logistic regression, the probability of success depends on more than one explanatory variable

$$p = \frac{e^{\alpha + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k}}{1 + e^{\alpha + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k}}$$

$$\text{and } \frac{p}{1-p} = e^{\alpha + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k}$$

- Therefore

$$\ln(p/(1-p)) = \alpha + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k$$

Interpretation of the coefficients

- Example: 2 independent variables, X_1 a numerical variable, X_2 a dichotomous exposure variable
- $\text{logit}(p) = \log(p/(1-p)) = \alpha + \beta_1 x_1 + \beta_2 x_2$
- When $x_2=1$ (exposed), holding x_1 constant,
$$\text{logit}(p_1) = \alpha + \beta_1 x_1 + \beta_2$$
- When $x_2=0$ (not exposed), holding x_1 constant,
$$\text{logit}(p_0) = \alpha + \beta_1 x_1$$

Interpretation of the coefficients

So $\text{logit}(p_1/p_0) = (\alpha + \beta_1 x_1 + \beta_2) - (\alpha + \beta_1 x_1) = \beta_2$

So
$$\log \left(\frac{\frac{p_1}{1-p_1}}{\frac{p_0}{1-p_0}} \right) = \beta_2$$

so $OR = e^{\beta_2}$ adjusted for x_1

```
. logistic coldany age_10 rested_mostly
```

Logistic regression	Number of obs	=	450
	LR chi2(2)	=	10.67
	Prob > chi2	=	0.0048
Log likelihood = -266.96282	Pseudo R2	=	0.0196

coldany	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
-----+-----						
age_10	.7774313	.0715981	-2.73	0.006	.6490378	.9312238
rested_mos~y	.7685218	.181155	-1.12	0.264	.4841857	1.219833

For this example the 0.777 is the odds ratio for a ten-year difference in age when you hold rested status constant

0.768 is the odds ratio for being well rested when you hold age constant

- Want more on linear regression and logistic regression? Go to

<http://www.ats.ucla.edu/stat/stata/webbooks/>

- Explore correlates of having a cold in our class data set. Hours of sleep? Having children under the age of 6? These while holding age constant?

Statistical hypothesis tests

Data and comparison type	Alternative hypotheses	Parametric test Stata command	Non-parametric test Stata command
Numerical; One mean	$H_a: \mu \neq \mu_a$ (two-sided) $H_a: \mu > \mu_a$ or $\mu < \mu_a$ (one-sided)	Z or t-test • <code>ttest var1=hypoth val.*</code>	
Numerical; Two means, <u>paired data</u>	$H_a: \mu_1 \neq \mu_2$ (two-sided) $H_a: \mu_1 > \mu_2$ or $\mu < \mu_a$ (one-sided)	Paired t-test • <code>ttest var1=var2*</code>	Sign test • <code>signtest var1=var2</code> • Wilcoxon Signed-Rank <code>signrank var1=var2</code>
Numerical; Two means, independent data	$H_a: \mu_1 \neq \mu_2$ (two-sided) $H_a: \mu_1 > \mu_2$ or $\mu < \mu_a$ (one-sided)	T-test (equal or unequal variance) • <code>ttest var1, by(byvar)</code> <code>unequal</code>	Wilcoxon rank-sum test • <code>ranksum var1, by(byvar)</code>
Numerical, Two or more means, independent data	$H_a: \mu_1 \neq \mu_2$ or $\mu_1 \neq \mu_3$ or $\mu_2 \neq \mu_3$ etc.	ANOVA • <code>oneway var1 byvar</code>	Kruskal Wallis test • <code>kwallis var1, by(byvar)</code>
Dichotomous; One proportion	$H_a: p \neq p_a$ (two-sided) $H_a: p > p_a$ or $p < p_a$ (one-sided)	Proportion test • <code>prtest var1=hypoth value*</code> • <code>bitest var1=hypoth value</code>	
Dichotomous; two proportions	$H_a: p_1 \neq p_2$ (two-sided) $H_a: p_1 > p_2$ (one-sided)	Proportion test (z-test) • <code>prtest var1, by(byvar)</code>	Chi-square test • <code>tab var1 var2, chi exact</code> McNemar's for paired data: • <code>mcc var1 var2</code>
Categorical by categorical (nxk)	H_a : The rows not independent of the columns		Chi-square test • <code>tab var1 var2, chi exact</code>

Further methods

Dependent variable (y)	Independent variable(s) type(s) (x)	Methods	Non-parametric methods
Numerical/continuous	One numerical/continuous variable	Pearson correlation; simple linear regression	Spearman rank correlation
Numerical/continuous	One categorical variable	Linear regression, t-test, ANOVA	Sign test, Signed rank test, Wilcoxon rank sum, Kruskal-Wallis
Numerical/continuous	Multiple variables/types of variables	Linear regression	
Dichotomous	One categorical variable	Logistic regression (or proportion test for 2x2)	Chi-square test or Fisher exact test
Dichotomous	One numerical/continuous variable	Logistic regression	
Dichotomous	Multiple variables/types of variables	Logistic regression	

Last sessions

- Thursday, Dec. 5 – no lab
- Final exam
 - Take home exam will be posted online on Thursday Dec. 5
 - The exam is worth 30% of your grade.
 - NO COLLABORATIONS WITH YOUR CLASSMATES OR OTHERS
 - Exam due by e-mail to TAs Thursday, December 12, 10:30 a.m.
- Tuesday Dec. 10 – office hours by appointment

Thank you and good luck!