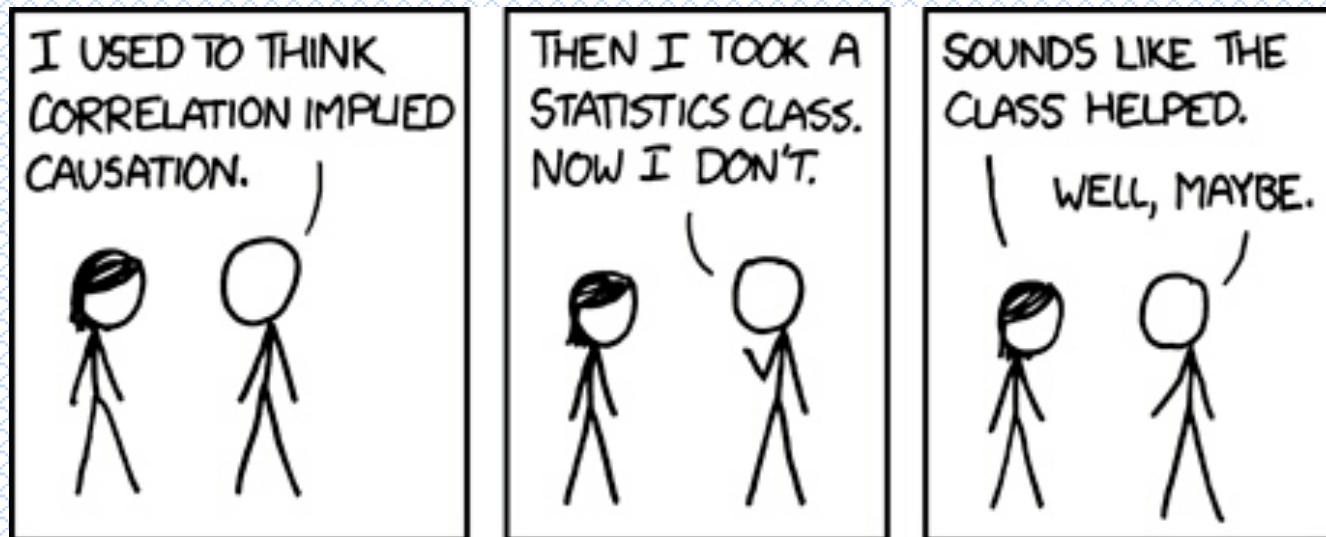


Biostat 200

Linear and Logistic Models



Correlation is not Causation

WHY IS THAT WOMAN SCOWLING
AT ME? DO I KNOW HER?



If she loves you more each and every day,
by linear regression she hated you before you met.

Review: Inference for **predicted values**

- We might want to estimate the mean value of y at a particular value of x
- What is the mean FEV for children who are 10 years old? $\hat{y} = .432 + .222 * x = .432 + .222 * 10 = 2.643$ liters
- We can construct a 95% confidence interval for the estimated mean
 - $(\hat{y} - t_{n-2, .025} se(\hat{y}), \hat{y} + t_{n-2, .025} se(\hat{y}))$, where

$$se(\hat{y}) = s_{y|x} \sqrt{\frac{1}{n} + \frac{(x - \bar{x})^2}{\sum_{i=1}^n (x_i - \bar{x})^2}}$$

$$\text{where } s_{y|x} = \sqrt{\frac{\sum_{i=1}^n (y_i - \hat{y})^2}{n - 2}} = \sqrt{\frac{RSS}{n - 2}}$$

Use Stata for your prediction

- Stata will calculate the fitted regression values and the standard errors
 - `regress fev age`
 - `predict fev_pred, xb` -> predicted mean values ($\hat{y}$)
 - `predict fev_predse, stdp` -> se of $\hat{y}$ values

New variable names that I made up

You don't have to calculate these to get a plot with the 95% CI:

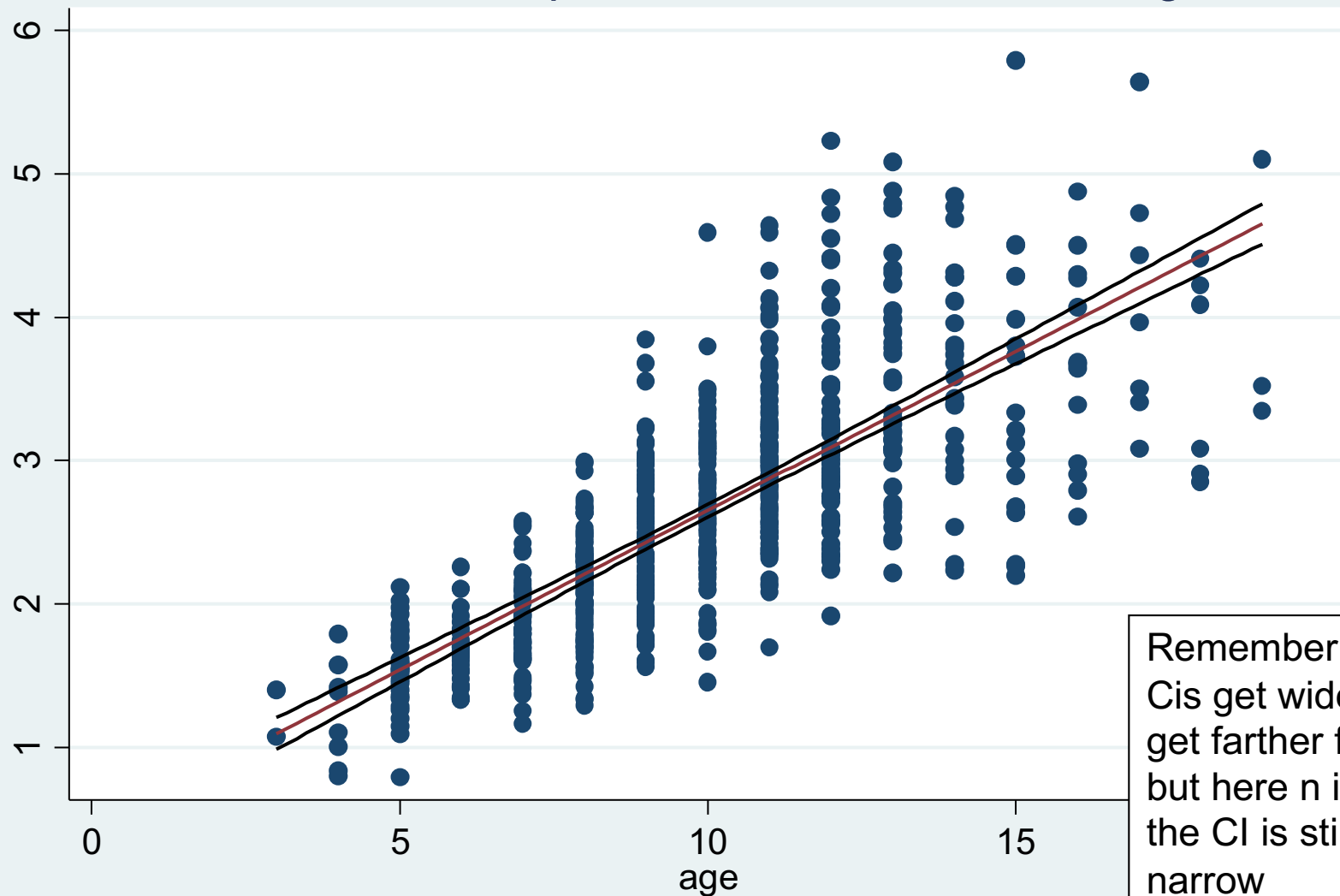
```
twoway (lfitci fev age)
```

Stands for linear fit for the line, CI for the confidence interval

```
. list fev age fev_pred fev_predse
```

	fev	age	fev_pred	fev_predse
1.	1.708	9	2.430017	.0232702
2.	1.724	8	2.207976	.0265199
3.	1.72	7	1.985935	.0312756
4.	1.558	9	2.430017	.0232702
5.	1.895	9	2.430017	.0232702
6.	2.336	8	2.207976	.0265199
7.	1.919	6	1.763894	.0369605
8.	1.415	6	1.763894	.0369605
9.	1.987	8	2.207976	.0265199
10.	1.942	9	2.430017	.0232702
11.	1.602	6	1.763894	.0369605
12.	1.735	8	2.207976	.0265199
13.	2.193	8	2.207976	.0265199
14.	2.118	8	2.207976	.0265199
15.	2.258	8	2.207976	.0265199
336.	3.147	13	3.318181	.0320131
337.	2.52	10	2.652058	.0221981
338.	2.292	10	2.652058	.0221981

95% CI for the predicted means for each age



Remember that the CIs get wider as you get farther from $\bar{x}$; but here n is large so the CI is still very narrow

```
twoway (scatter fev age) (lfitci fev age, ciplot(rline)
blcolor(black)), legend(off) title(95% CI for the predicted
means for each age )
```

Prediction intervals

- The intervals we just made were for means of y at particular values of x
- What if we want to predict the FEV value for an *individual child* at age 10?
- The prediction is the same – plug into the regression equation: $\tilde{y} = .432 + .222 * 10 = 2.643$ liters
- But the standard error of $\tilde{y}$ is not the same as the standard error of $\hat{y}$

Prediction intervals

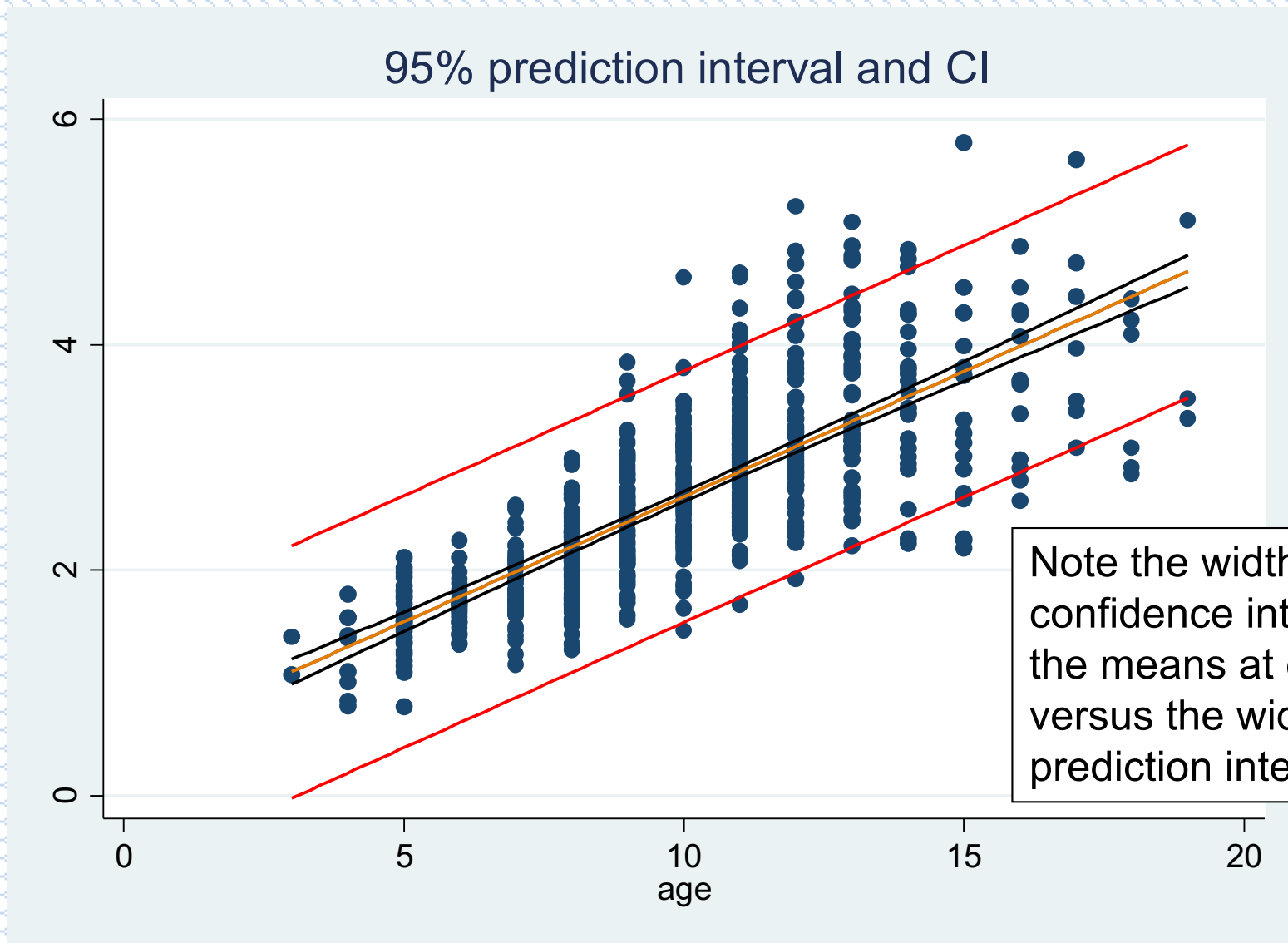
$$\begin{aligned} s\hat{e}(\tilde{y}) &= s_{y|x} \sqrt{1 + \frac{1}{n} + \frac{(x - \bar{x})^2}{\sum_{i=1}^n (x_i - \bar{x})^2}} \\ &= \sqrt{s_{y|x}^2 + \frac{s_{y|x}^2}{n} + \frac{s_{y|x}^2 (x - \bar{x})^2}{\sum_{i=1}^n (x_i - \bar{x})^2}} \end{aligned}$$

- This differs from the $se(\hat{y})$ only by the extra variance of y in the formula but it makes a big difference
- There is much more uncertainty in predicting a future value versus predicting a mean
- Stata will calculate these using `predict fev_predse_ind, stdf`

Where the **f** is for forecast

```
. list fev age fev_pred fev_predse fev_pred_ind
```

	fev	age	fev_pred	fev~edse	fev~ndse
1.	1.708	9	2.430017	.0232702	.5680039
2.	1.724	8	2.207976	.0265199	.5681463
3.	1.72	7	1.985935	.0312756	.5683882
4.	1.558	9	2.430017	.0232702	.5680039
5.	1.895	9	2.430017	.0232702	.5680039
6.	2.336	8	2.207976	.0265199	.5681463
7.	1.919	6	1.763894	.0369605	.5687293
8.	1.415	6	1.763894	.0369605	.5687293
9.	1.987	8	2.207976	.0265199	.5681463
10.	1.942	9	2.430017	.0232702	.5680039
11.	1.602	6	1.763894	.0369605	.5687293
12.	1.735	8	2.207976	.0265199	.5681463
13.	2.193	8	2.207976	.0265199	.5681463
14.	2.118	8	2.207976	.0265199	.5681463
15.	2.258	8	2.207976	.0265199	.5681463
336.	3.147	13	3.318181	.0320131	.5684292
337.	2.52	10	2.652058	.0221981	.567961
338.	2.292	10	2.652058	.0221981	.567961



```

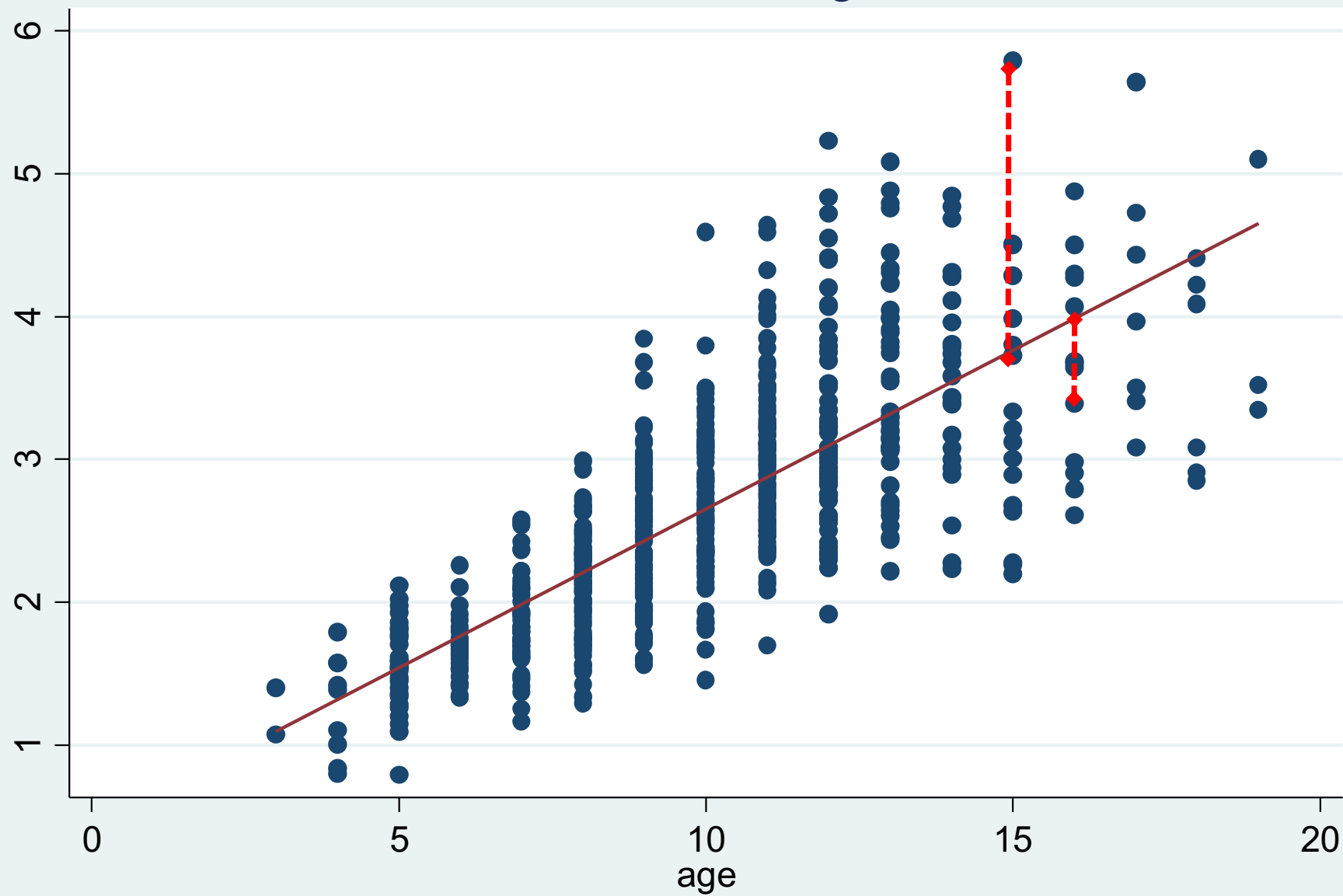
twoway (scatter fev age)      (lfitci fev age, ciplot(rline)
blcolor(black) ) (lfitci fev age, stdf ciplot(rline)
blcolor(red) ), legend(off) title(95% prediction interval and
CI )

```

Model fit -- Residuals

- Residuals are the difference between the observed y values and the regression line for each value of x : $y_i - \hat{y}_i$
- If all the points lie along a straight line, the residuals are all 0
- If there is a lot of variability at each level of x , the residuals are large
- The sum of the squared residuals is what was minimized in the least squares method of fitting the line

FEV versus age



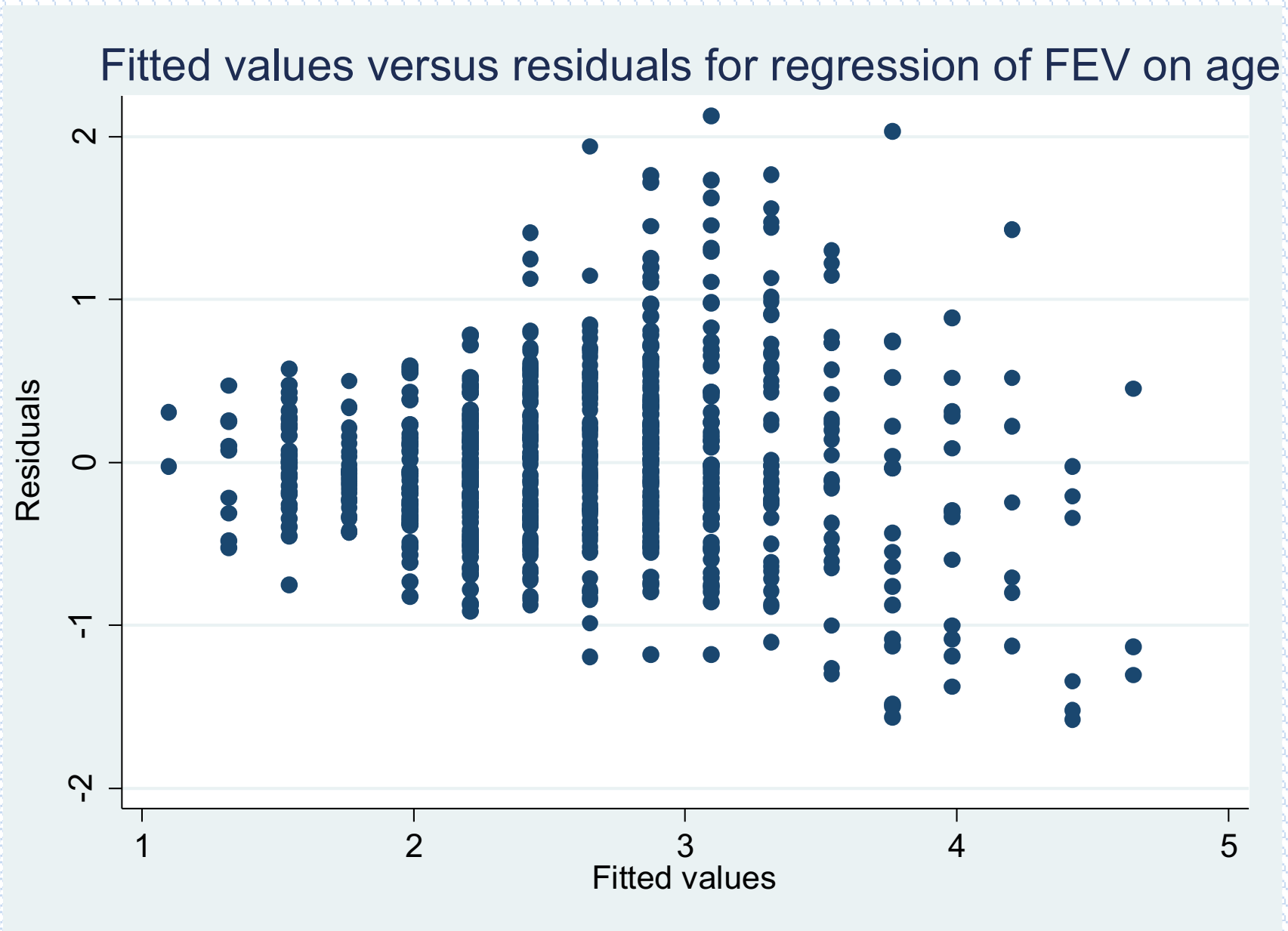
Residuals

- We examine the residuals using scatter plots
- We plot the fitted values $\hat{y}_i$ on the x-axis and the residuals $y_i - \hat{y}_i$ on the y-axis
- We use the fitted values because they have the effect of the independent variable removed
- To calculate the residuals and the fitted values

Stata:

```
regress fev age  
rvfplot
```

- RVF = residuals versus fitted



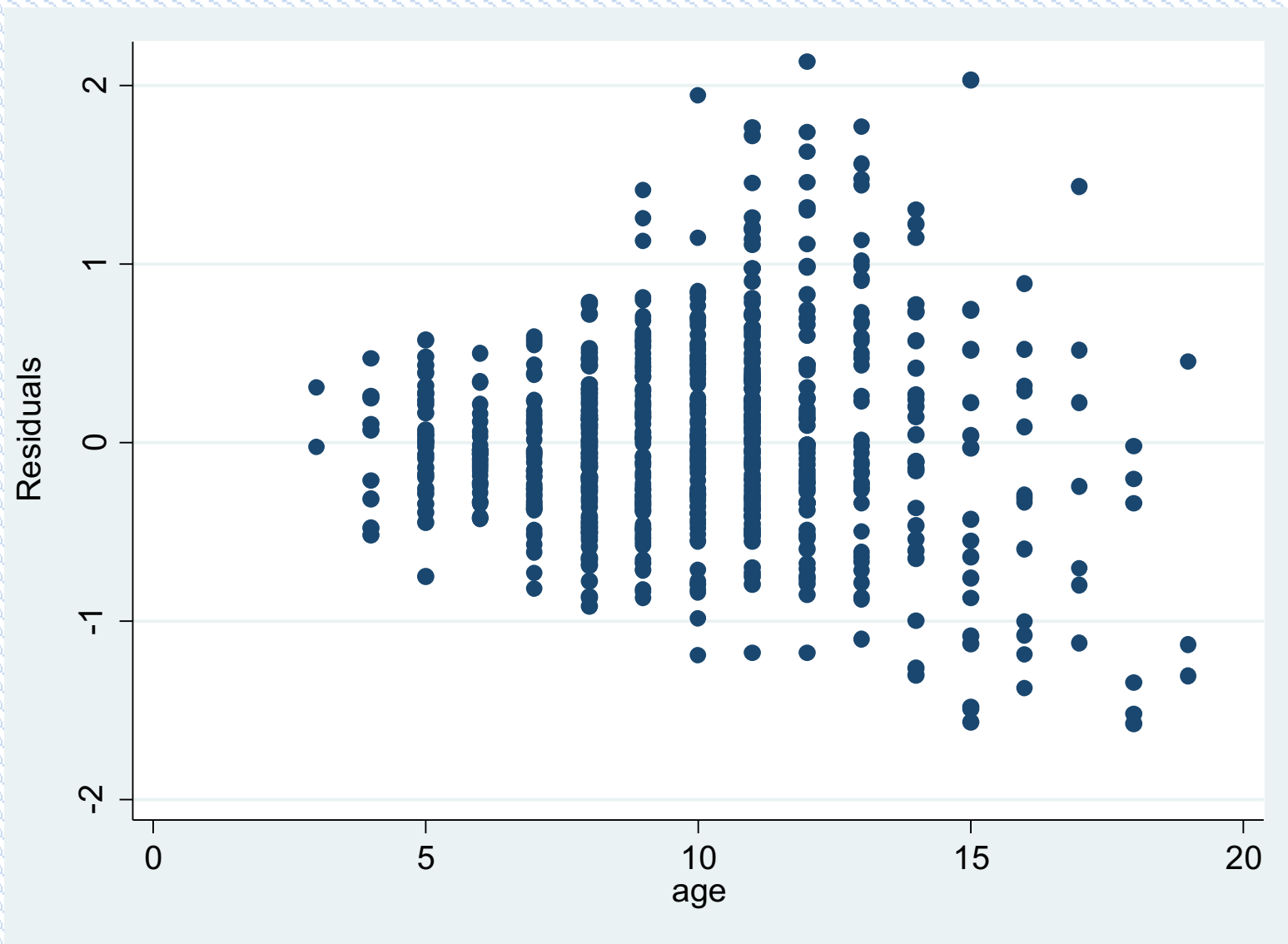
```
rvfplot, title(Fitted values versus residuals for  
regression of FEV on age)
```

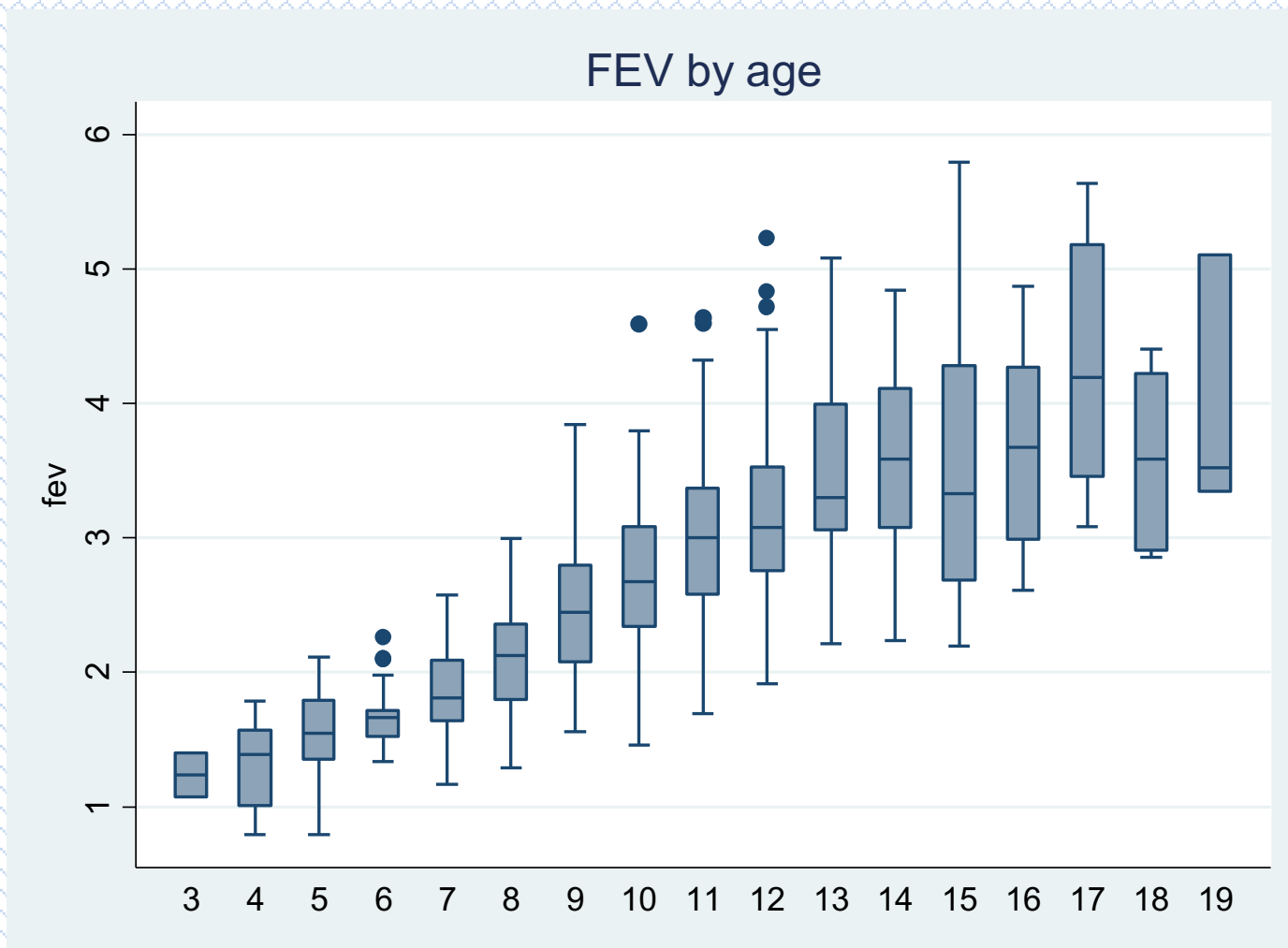
Residuals

- This plot shows that as the fitted value of FEV increases, the spread of the residuals increase – this suggests heteroscedasticity
- We would get a similar plot if we plotted age on the x-axis

```
rvpplot age, name(res_v_age)
```

- **RVP=residuals versus predictor**
- We had a hint of this when looking at the box plots of FEV by age groups in the previous lecture





`graph box fev, over(age) title(FEV by age)`

Note that heteroscedasticity does not bias the estimates of the parameters, but it does reduce the precision of the estimates (and therefore reduces power)

Transformations

- One way to deal with this is to transform either x or y or both
- A common transformation is the log transformation
- Log transformations bring large values closer to the rest of the data
- There are methods to correct the standard errors for heteroscedasticity other than transformations

Log transformations

Value	Ln	Log ₁₀
0	$-\infty$	$-\infty$
0.001	-6.91	-3.00
0.05	-3.00	-1.30
1	0.00	0.00
5	1.61	0.70
10	2.30	1.00
50	3.91	1.70
103	4.63	2.01

Be careful of $\log(0)$ or $\ln(0)$

Be sure you know which log base your computer program is using

- In Stata $\log()$ will give you $\ln()$
- If you want $\log_{10}$ use $\log_{10}()$

Transformations

- Transform FEV to $\ln(\text{FEV})$

```
. gen fev_ln=log(fev)
. summ fev fev_ln
```

Variable	Obs	Mean	Std. Dev.	Min	Max
-----+-----					
fev	654	2.63678	.8670591	.791	5.793
fev_ln	654	.915437	.3332652	-.2344573	1.75665

- Examine the histograms
- Run the regression of $\ln(\text{FEV})$ on age and examine the residuals

```
regress fev_ln age
rvfplot, title(Fitted values versus residuals for
               regression of  $\ln\text{FEV}$  on age)
```

Regressions: fev & fev_ln

```
. regress fev_ln age
```

Source	SS	df	MS	Number of obs	=	654
Model	43.2100544	1	43.2100544	F(1, 652)	=	961.01
Residual	29.3158601	652	.044962976	Prob > F	=	0.0000
Total	72.5259145	653	.111065719	R-squared	=	0.5958
				Adj R-squared	=	0.5952
				Root MSE	=	.21204

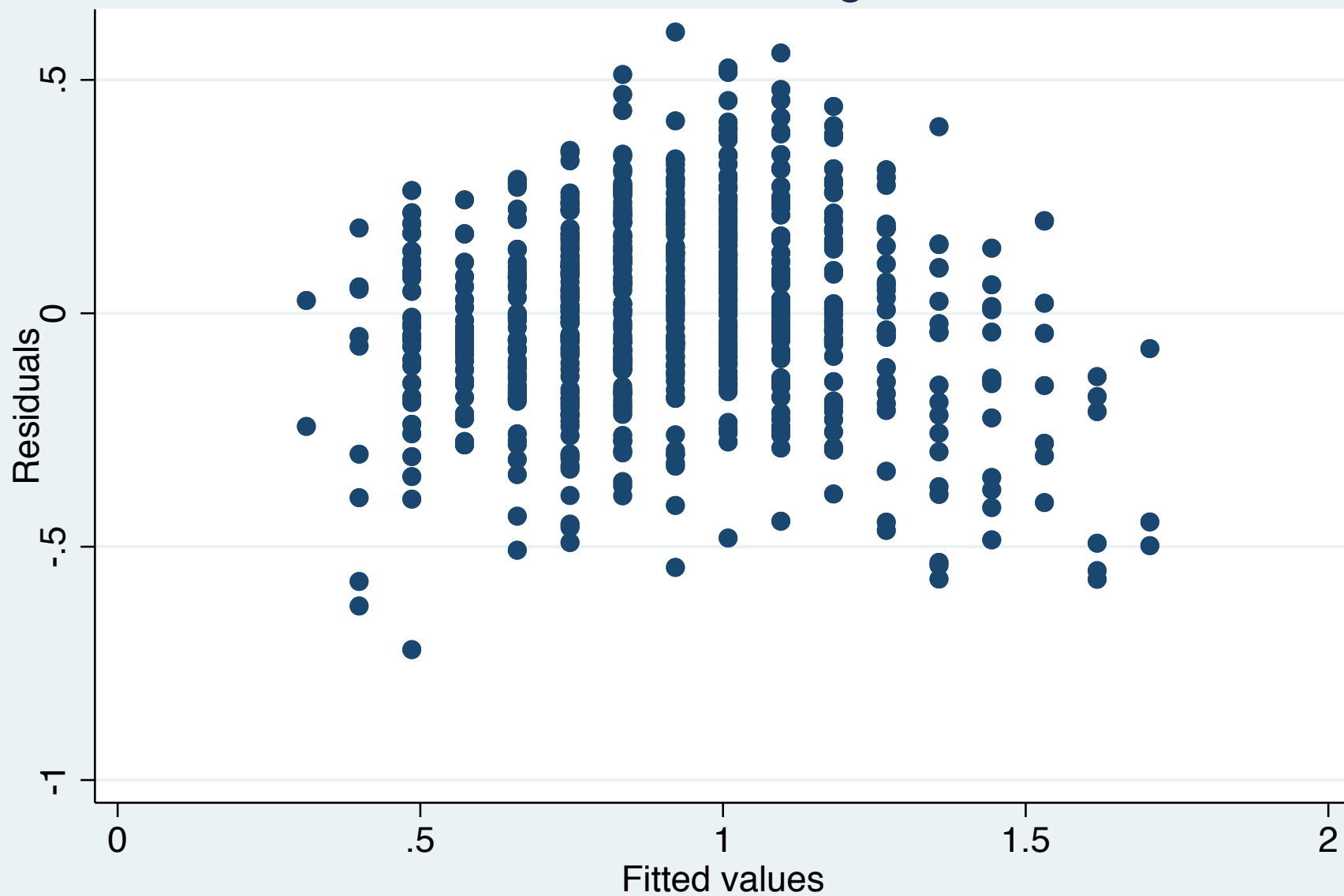
fev_ln	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	.0870833	.0028091	31.00	0.000	.0815673	.0925993
_cons	.050596	.029104	1.74	0.083	-.0065529	.1077449

```
. regress fev age
```

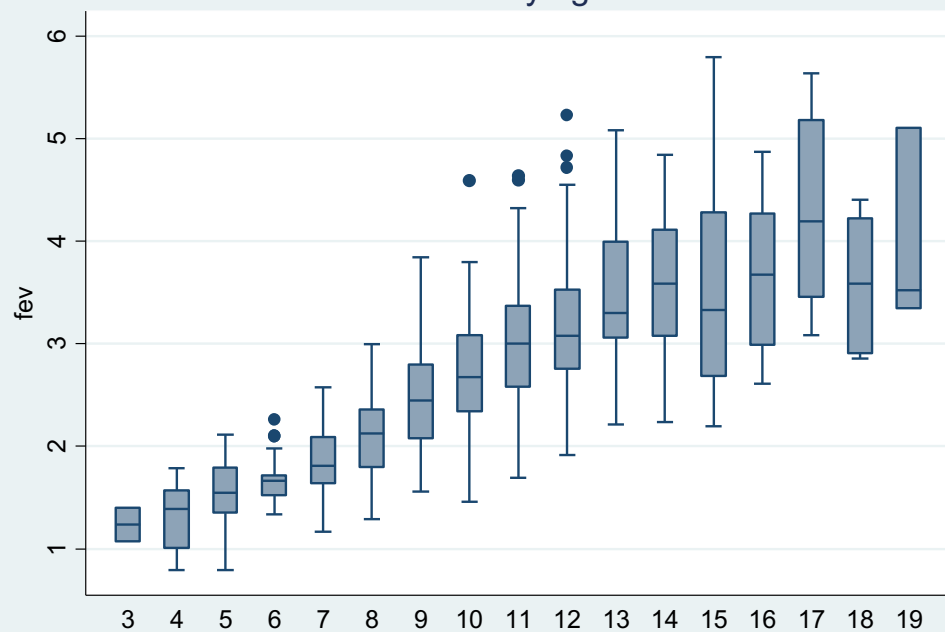
Source	SS	df	MS	Number of obs	=	654
Model	280.919154	1	280.919154	F(1, 652)	=	872.18
Residual	210.000679	652	.322086931	Prob > F	=	0.0000
Total	490.919833	653	.751791475	R-squared	=	0.5722
				Adj R-squared	=	0.5716
				Root MSE	=	.56753

fev	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	.222041	.0075185	29.53	0.000	.2072777	.2368043
_cons	.4316481	.0778954	5.54	0.000	.278692	.5846042

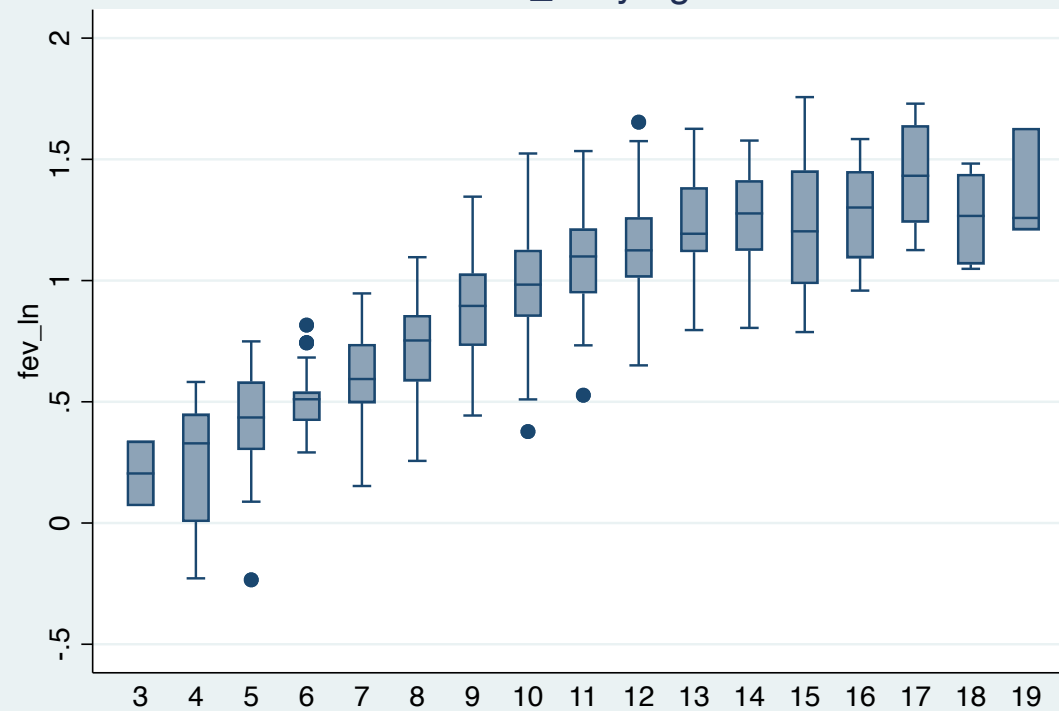
Fitted values versus residuals for regression of lnFEV on age



FEV by age



FEV_In by age



Interpretation of regression coefficients for transformed y value

- The regression equation is:

$$\ln(FEV) = \hat{\alpha} + \hat{\beta} * age$$
$$= 0.051 + 0.087 \text{ age}$$

- So a one year change in age corresponds to a .087 change in $\ln(FEV)$
- The change is on a multiplicative scale, so if you exponentiate, you get a percent change in y
- $e^{0.087} = 1.091$ or a 9.1% change in FEV

For example...

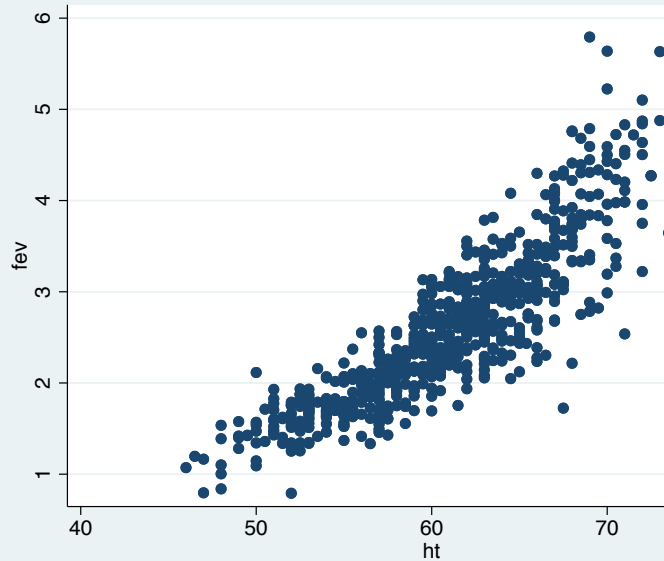
- $\ln(\text{FEV}) = 0.051 + 0.087 \text{ age}$
- $\ln(\text{FEV}_{\text{age21}}) = 0.051 + 0.087 * 21$
- $\ln(\text{FEV}_{\text{age20}}) = 0.051 + 0.087 * 20$
- $\ln(\text{FEV}_{\text{age21}}) - \ln(\text{FEV}_{\text{age20}}) = 0.087$

Remember $\ln(a) - \ln(b) = \ln(a/b)$

- $\ln(\text{FEV}_{\text{age21}} / \text{FEV}_{\text{age20}}) = 0.087$
- $\text{FEV}_{\text{age21}} / \text{FEV}_{\text{age20}} = e^{0.087} = 1.09$

Now using height as the independent variable

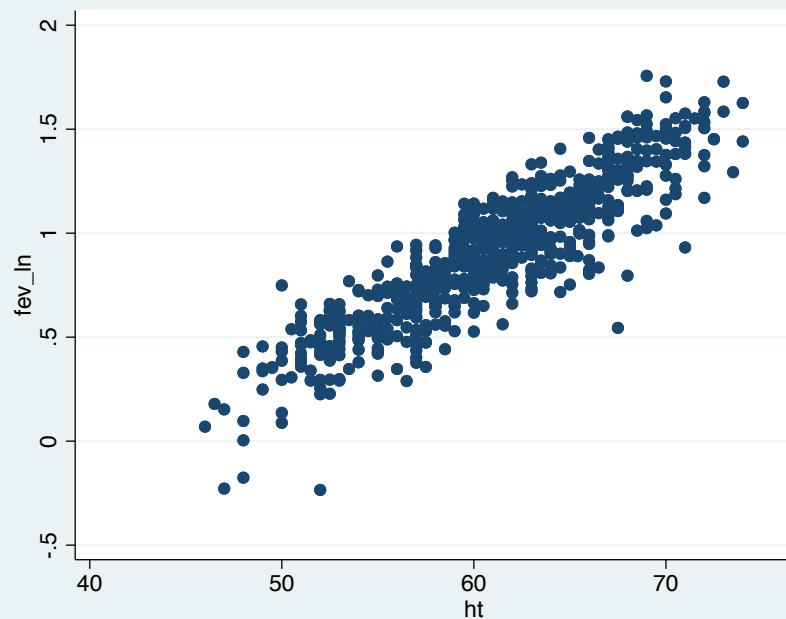
1. Make a scatter plot of FEV by height and FEV_In by height
2. Run a regression of FEV on height and FEV_In on height
3. What does this tell you?



```
. regress fev ht
```

Source	SS	df	MS	Number of obs	=	654
Model	369.985854	1	369.985854	F(1, 652)	=	1994.73
Residual	120.933979	652	.185481563	Prob > F	=	0.0000
				R-squared	=	0.7537
				Adj R-squared	=	0.7533
Total	490.919833	653	.751791475	Root MSE	=	.43068

fev	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
ht	.1319756	.002955	44.66	0.000	.1261732	.137778
_cons	-5.432679	.1814599	-29.94	0.000	-5.788995	-5.076363



```
. regress fev_ln ht
```

Source	SS	df	MS	Number of obs	=	654
Model	57.7021467	1	57.7021467	F(1, 652)	=	2537.94
Residual	14.8237679	652	.02273584	Prob > F	=	0.0000
				R-squared	=	0.7956
				Adj R-squared	=	0.7953
Total	72.5259145	653	.111065719	Root MSE	=	.15078

fev_ln	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
ht	.0521191	.0010346	50.38	0.000	.0500876	.0541506
_cons	-2.271312	.063531	-35.75	0.000	-2.396062	-2.146562

Categorical independent variables

- We previously noted that the independent variable (the X variable) does not need to be normally distributed
- In fact, this variable can be categorical

Categorical independent variables

- Coding: Dichotomous variables in regression models
 - 1 to represent the level of interest
 - 0 to represent the comparison or reference group.These 0-1 variables are called indicator or dummy variables.
- The regression model is the same
- The interpretation is slightly different
 - β is the change in y that corresponds to being in the group of interest vs. the reference category

Categorical independent variables

- Example sex: female $x_{\text{sex}}=0$, for male $x_{\text{sex}}=1$

- Regression of FEV and sex

- $$\hat{\text{fev}} = \hat{\alpha} + \hat{\beta} x_{\text{sex}}$$

- For male: $\hat{\text{fev}}_{\text{male}} = \hat{\alpha} + \hat{\beta} * 1 = \hat{\alpha} + \hat{\beta}$

- For female: $\hat{\text{fev}}_{\text{female}} = \hat{\alpha} + \hat{\beta} * 0 = \hat{\alpha}$

$$\text{So } \hat{\text{fev}}_{\text{male}} - \hat{\text{fev}}_{\text{female}} = \hat{\alpha} + \hat{\beta} - \hat{\alpha} = \hat{\beta}$$

- Remember, $\hat{\alpha}$ is the mean value of y when $x=0$

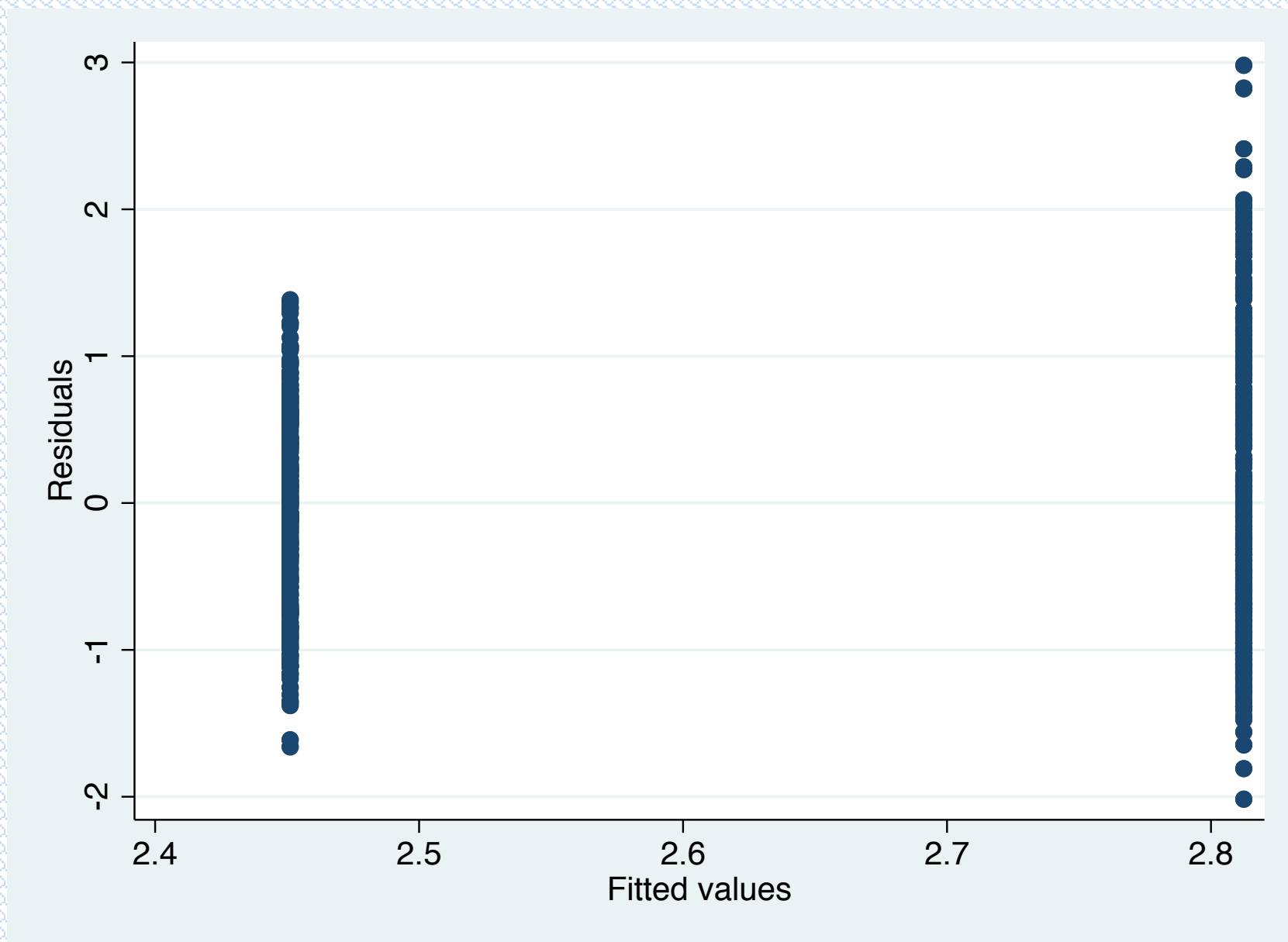
So here it is the mean FEV for sex=female

```
. regress fev sex
```

Source	SS	df	MS	Number of obs	=	654
Model	21.3239848	1	21.3239848	F(1, 652)	=	29.61
Residual	469.595849	652	.720239032	Prob > F	=	0.0000
				R-squared	=	0.0434
				Adj R-squared	=	0.0420
Total	490.919833	653	.751791475	Root MSE	=	.84867

fev	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
sex	.3612766	.0663963	5.44	0.000	.2309002	.491653
_cons	2.45117	.047591	51.50	0.000	2.35772	2.54462

1. What is the estimate for beta? How is it interpreted?
2. What is the estimate for alpha? How is it interpreted?
3. What hypothesis is tested where it says P>|t|?
4. How much of the variance in FEV is explained by sex?
5. What does the rvfplot look like?



Categorical independent variable

- Remember that the regression equation is

$$\mu_{y/x} = \alpha + \beta x$$

- The only variables x can take are 0 and $\mu_{y/0} = \alpha$
 $\mu_{y/1} = \alpha + \beta$
- The estimated mean FEV for females is $\hat{\alpha}$ and the estimated mean FEV for males is $\hat{\alpha} + \hat{\beta}$
- What other test have we learned that tests the same thing?

Two-sample t test with equal variances

Ha: diff < 0	Ha: diff != 0	Ha: diff > 0
Pr(T < t) = 0.0000	Pr(T > t) = 0.0000	Pr(T > t) = 1.0000

Source	SS	df	MS	Number of obs	=	654
				F(1, 652)	=	29.61
Model	21.3239848	1	21.3239848	Prob > F	=	0.0000
Residual	469.595849	652	.720239032	R-squared	=	0.0434
				Adj R-squared	=	0.0420
Total	490.919833	653	.751791475	Root MSE	=	.84867

35

Categorical independent variables

- e.g. Race group = White, Asian/PI, Other
- If Race=White is set as reference category, dummy variables look like:

	$X_{\text{Asian/PI}}$	X_{Other}
White	0	0
Asian/PI	1	0
Other	0	1

Categorical independent variables

- Then the regression equation is:

$$y = \alpha + \beta_1 x_{\text{Asian/PI}} + \beta_2 x_{\text{Other}} + \varepsilon$$

- For race group=White

$$\hat{y} = \hat{\alpha} + \hat{\beta}_1 0 + \hat{\beta}_2 0 = \hat{\alpha}$$

- For race group=Asian/PI

$$\hat{y} = \hat{\alpha} + \hat{\beta}_1 1 + \hat{\beta}_2 0 = \hat{\alpha} + \hat{\beta}_1$$

- For race group=other

$$\hat{y} = \hat{\alpha} + \hat{\beta}_1 0 + \hat{\beta}_2 1 = \hat{\alpha} + \hat{\beta}_2$$

Stata indicator variables

- You actually don't have to make the dummy variables yourself
- All you have to do is tell Stata that a variable is categorical using `i.` before a variable name
- The regression of FEV regressed on age group

```
regress fev i.agecat
```

```
. regress fev i.agecat
```

Source	SS	df	MS	Number of obs = 654		
Model	265.481813	4	66.3704533	F(4, 649) = 191.07		
Residual	225.43802	649	.347362127	Prob > F = 0.0000		
Total	490.919833	653	.751791475	R-squared = 0.5408		
				Adj R-squared = 0.5380		
				Root MSE = .58937		

fev	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
agecat						
6-	.4713427	.104309	4.52	0.000	.2665188	.6761665
9-	1.244846	.1010818	12.32	0.000	1.046359	1.443332
12-	1.912191	.1081	17.69	0.000	1.699923	2.124459
15-	2.237758	.1264743	17.69	0.000	1.98941	2.486106
_cons	1.472385	.0943754	15.60	0.000	1.287067	1.657703

Multiple regression

- Additional explanatory variables might increase our understanding of a dependent variable
- We can posit the population equation

$$\mu_{y|x_1, x_2, \dots, x_q} = \alpha + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_q x_q$$

- α is the mean of y when all the explanatory variables are 0
- β_i is the change in the mean value of y that corresponds to a 1 unit change in x_i when all the other explanatory variables are held constant

Multiple regression

- $y = \alpha + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_q x_q + \varepsilon$
- Assumptions
 - $x_1, x_2, \dots, x_q$ are measured without error
 - The distribution of y is normal with mean $\mu_{y/x_1, x_2, \dots, x_q}$ and standard deviation $\sigma_{y/x_1, x_2, \dots, x_q}$
 - The population regression model holds
 - For any set of values of the explanatory variables, $x_1, x_2, \dots, x_q$, $\sigma_{y/x_1, x_2, \dots, x_q}$ is constant – homoscedasticity
 - The y outcomes are independent

Multiple regression – Least Squares

- We estimate the regression line

$$\hat{y} = \hat{\alpha} + \hat{\beta}_1 x_1 + \hat{\beta}_2 x_2 + \dots \hat{\beta}_q x_q$$

using the method of least squares to minimize

$$\begin{aligned} \sum_{i=1}^n e_i^2 &= \sum_{i=1}^n (y_i - \hat{y}_i)^2 \\ &= \sum_{i=1}^n [y_i - (\hat{\alpha} + \hat{\beta}_1 x_{1i} + \hat{\beta}_2 x_{2i} + \dots + \hat{\beta}_q x_{qi})]^2 \end{aligned}$$

Multiple regression

- For one explanatory variable – the regression model represents a straight line through a cloud of points -- in 2 dimensions
- With 2 explanatory variables, the model is a plane in 3 dimensional space (one for each variable)
- etc.
- In Stata we just add explanatory variables to the regress statement
- `regress fev age ht`

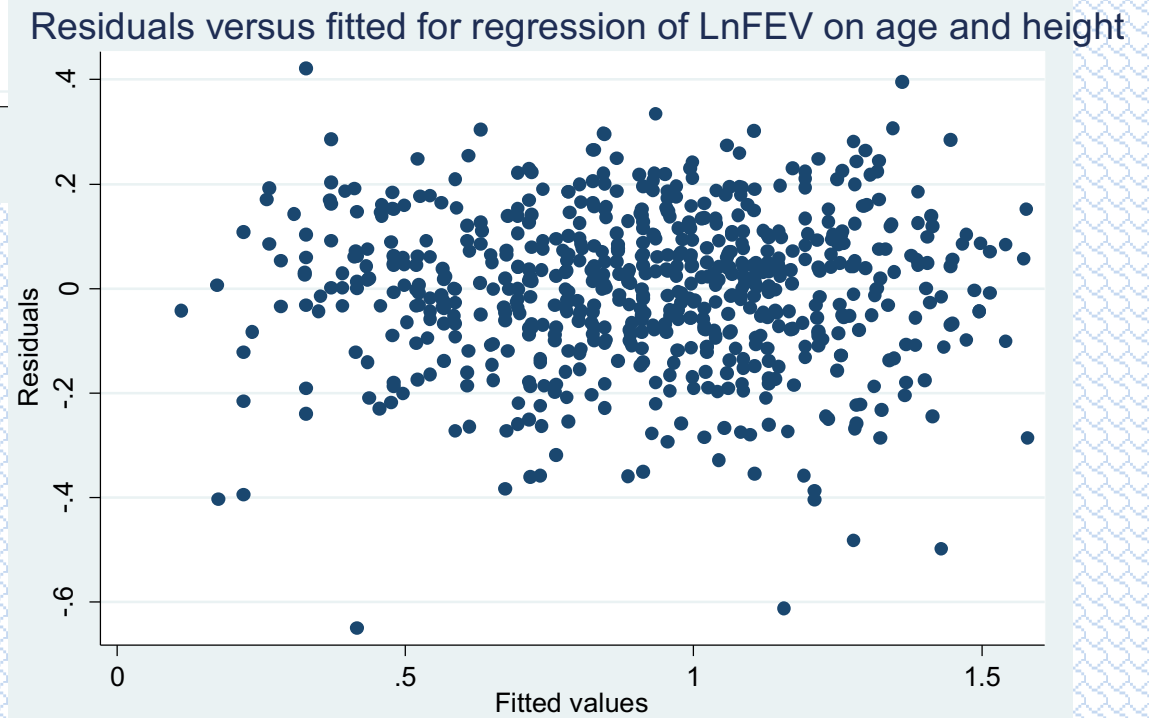
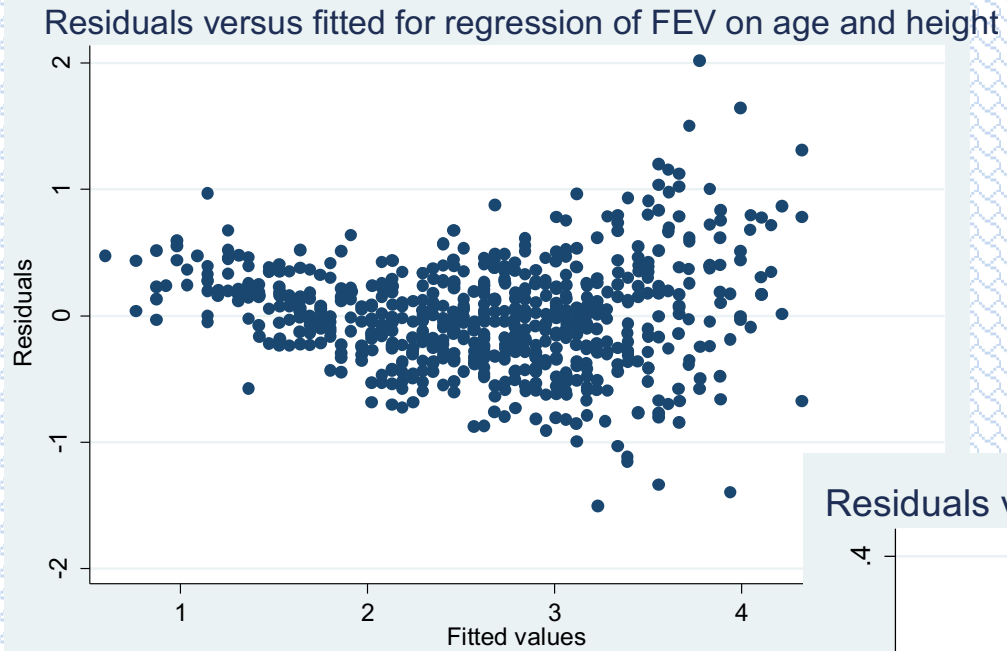
```
. regress fev age ht
```

Source	SS	df	MS	Number of obs = 654		
Model	376.244941	2	188.122471	F(2, 651)	=	1067.96
Residual	114.674892	651	.176151908	Prob > F	=	0.0000
Total	490.919833	653	.751791475	R-squared	=	0.7664
				Adj R-squared	=	0.7657
				Root MSE	=	.4197

fev	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	.0542807	.0091061	5.96	0.000	.0363998	.0721616
ht	.1097118	.0047162	23.26	0.000	.100451	.1189726
_cons	-4.610466	.2242706	-20.56	0.000	-5.050847	-4.170085

- Now the F-test has 2 degrees of freedom in the numerator because there are 2 explanatory variables
- R^2 will always increase as you add more variables into the model
- The Adj R-squared accounts for the addition of variables and is comparable across models with different numbers of parameters

Examine the residuals...



Logistic regression

- For linear regression
 - The predicted values can be from $-\infty$ to $+\infty$
 - Y values are assumed to follow a normal distribution
- For many situations we want to model a dichotomous outcome, such as disease/no disease and cannot use linear regression
- We want to model p , the probability of disease

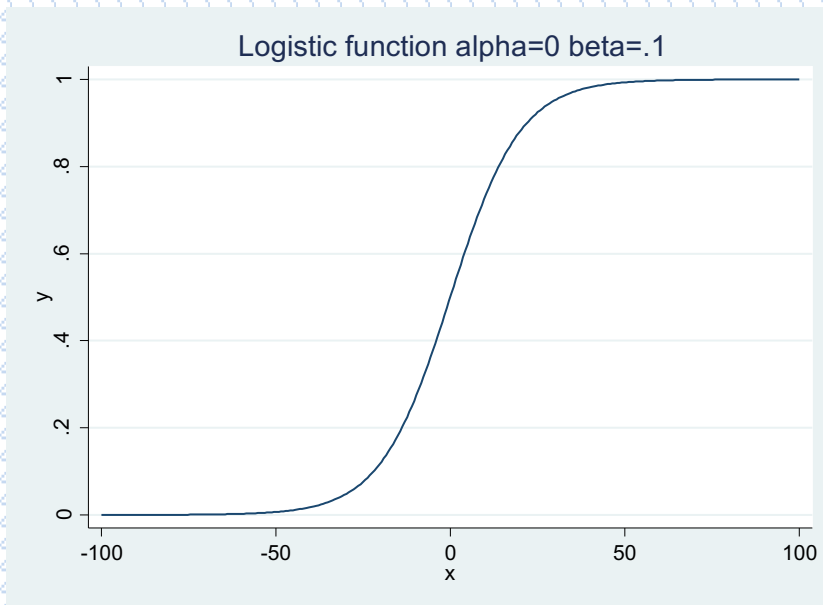
Logistic regression

- $p = P(Y=1)$ i.e. $P(\text{outcome}=\text{disease})$
- A model of the form $p = \alpha + \beta x$ would be able to take on negative values or values more than 1
- $p = e^{\alpha + \beta x}$ is an improvement because it cannot be negative, but it still could be greater than 1

Logistic regression

- How about the function?

$$p = \frac{e^{\alpha + \beta x}}{1 + e^{\alpha + \beta x}}$$



- When $\alpha + \beta x = 0$, $p = 1/(1+1) = 0.5$
- When x =big negative number, p nears $0/(1+0) = 0$
- When x =big positive number, p nears $(\text{big number})/(1+\text{big number}) = 1$
- The function models the probability slowly increasing over the values of X , until there is a steep rise, and another leveling off

Logistic regression

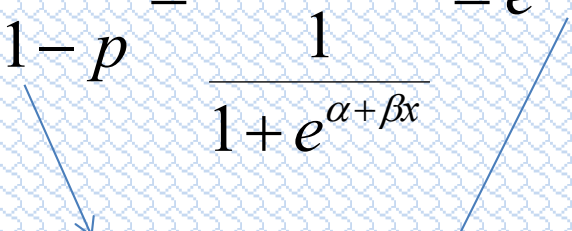
- The logistic function

$$p = \frac{e^{\alpha + \beta x}}{1 + e^{\alpha + \beta x}}$$

- Now check this out:

$$1 - p = 1 - \frac{e^{\alpha + \beta x}}{1 + e^{\alpha + \beta x}} = \frac{1 + e^{\alpha + \beta x} - e^{\alpha + \beta x}}{1 + e^{\alpha + \beta x}} = \frac{1}{1 + e^{\alpha + \beta x}}$$

- So odds of disease/success are

$$\frac{p}{1-p} = \frac{\frac{e^{\alpha + \beta x}}{1 + e^{\alpha + \beta x}}}{\frac{1}{1 + e^{\alpha + \beta x}}} = e^{\alpha + \beta x}$$


- And therefore $\ln(p/(1-p)) = \ln(e^{\alpha + \beta x}) = \alpha + \beta x$

Logistic regression

- So instead of assuming that the relationship between p and X is linear (as in linear regression), we are assuming that the relationship between $\ln(p/(1-p))$ and X is linear.
- $\ln(p/(1-p))$ is called the logit function
- $\text{Logit}(p) = \ln(p/(1-p))$
- In generalized linear models, some function of the outcome variable is assumed to be linear

e.g. $f(y) = \alpha + \beta x$ or $f(p) = \alpha + \beta x$

e.g. $\text{logit}(p) = \alpha + \beta x$

Logistic regression

- Assumptions
 - Y_i follows a bernoulli distribution (0,1)
 - The mean of Y at every X , $P(Y=1|x)$ is given by the logistic function
 - The values of the outcomes are independent
- Estimation
 - The parameters are estimated via a procedure called *maximum likelihood*
 - There is not a closed form solution of the maximization (i.e. we cannot write down a solution) – it is an iterative procedure

Interpretation of coefficients

- One dichotomous explanatory variable
- For $x=0$
 $\text{logit}(p_0) = \ln(p_0/(1-p_0))$
 $= \ln(P(Y=1 | X=0)/(1 - P(Y=1 | X=0)))$
 $= \ln(\text{odds of the outcome when } x=0)$
 $= \alpha + \beta * x = \alpha + \beta * 0 = \alpha$
- For $x=1$
 $\text{logit}(p_1) = \ln(p_1/(1-p_1))$
 $= \ln(P(Y=1 | X=1)/(1 - P(Y=1 | X=1)))$
 $= \text{odds of the outcome when } x=1$
 $= \alpha + \beta * 1 = \alpha + \beta$

Difference between treatment arms in PEth outcome (>50 ng/ml)*

```
logistic peth_audit_50 i.studyarm_n
```

Logistic regression

Number of obs = 324

LR chi2(1) = 1.14

Prob > chi2 = 0.2858

Log likelihood = -223.91123

Pseudo R2 = 0.0025

```
-----+-----  
peth_audit_50 | Odds Ratio   Std. Err.      z    P>|z|     [95% Conf. Interval]  
-----+-----  
1.studyarm_n |   .7869577   .1768227    -1.07   0.286    .5066332    1.222388  
    _cons |   1.05618   .1562084     0.37   0.712    .7903976    1.411335  
-----+-----
```

*BREATH cohort and comparisons_v3.dta

Difference between treatment arms in PEth outcome (>50 ng/ml)

logistic depvar indepvar

logistic peth_audit_50 i.studyarm_n

Logistic regression

Number of obs = 324
LR chi2(1) = 1.14
Prob > chi2 = 0.2858
Pseudo R2 = 0.0025

Log likelihood = -223.91123

peth_audit_50	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	

1.studyarm_n	.7869577	.1768227	-1.07	0.286	.5066332	1.222388
_cons	1.05618	.1562084	0.37	0.712	.7903976	1.411335

$=e^{\beta}$

This is used in comparing nested models of the same data set

Test of whether the model fit differs from the model with only α and no covariates

The Pseudo R^2 is an attempt to be analogous with linear regression but it is not

Stata logistic notes

- logistic depvar indepvars
 - gives you odds ratios
- logistic depvar indepvars, coef
 - gives you coefficients (alphas and betas)
- logit depvar indepvars
 - gives you coefficients
- logit depvar indepvars, or
 - gives you odds ratios

```
logit peth_50 i.studyarm_n
```

```
Iteration 0:    log likelihood = -198.01065
```

```
Iteration 1:    log likelihood = -197.99359
```

```
Iteration 2:    log likelihood = -197.99359
```

```
Logistic regression
```

```
Number of obs    =          287
```

```
LR chi2(1)       =           0.03
```

```
Prob > chi2      =          0.8534
```

```
Log likelihood = -197.99359
```

```
Pseudo R2       =          0.0001
```

```
-----+-----  
      peth_50 |      Coef.   Std. Err.      z    P>|z|     [95% Conf. Interval]  
-----+-----  
1.studyarm_n |   -.0453478   .2455377    -0.18   0.853    -.5265928    .4358972  
      _cons |   -.1438942   .1490438    -0.97   0.334    -.4360146    .1482262  
-----+-----
```

Odds ratio of PEth $\geq$ 50 for comparison vs. standard group

```
tabodds peth_audit_50 studyarm_n, or
```

studyarm_n	Odds Ratio	chi2	P>chi2	[95% Conf. Interval]	
0	1.000000	.	.	.	.
1	0.786958	1.13	0.2867	0.505895	1.224172
Test of homogeneity (equal odds):					
		chi2(1)	=	1.13	
		Pr>chi2	=	0.2867	
Score test for trend of odds:					
		chi2(1)	=	1.13	
		Pr>chi2	=	0.2867	

For one independent variable that is categorical, the odds ratio estimate is the same for logistic regression as for tabular methods.

What is the *odds ratio*?

- The odds ratio represents the change in the odds of membership in the “target” category (e.g., the category coded as 1.0), given a 1-unit increase in the predictor variable.
- Odds ratio is the increase (or decrease if negative) in odds of being in one outcome category when the value of the predictor variable increases by one unit.
- The odds ratio is calculated by the formula: e^g (where $e = 2.718$, and $g =$ the b coefficient from the logistic regression output).

What is the odds ratio?

- For example:
 - An odds ratio of 1.5 indicates that that outcome of “1” is 1.5 times as likely (or 50% more likely) with a 1-unit increase in the predictor. The odds increase by 50%.
 - An odds ratio of 0.80 indicates that an outcome of “1” is 0.8 times as likely (or 20% less likely [$1 - 0.8 = 0.2$]) with a 1-unit increase in the predictor; the odds are decreased by 20%.

How do you compute the probability of being in the “target” group, for a specific individual?

- Multiply each predictor value by its “b” coefficient, and add them to the constant. That sum is called “g.”
- Use “g” in the following formula:
- $e^g / 1 + e^g$

Continuous explanatory variable

```
. . logistic peth_audit_50 age_b
```

Logistic regression

Number of obs = 324
LR chi2(1) = 5.32
Prob > chi2 = 0.0210
Pseudo R2 = 0.0119

Log likelihood = -221.81862

peth_audit_50	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]		
-----+-----							
age_b	1.029989	.0133876	2.27	0.023	1.004081	1.056565	
_cons	.3782632	.1590054	-2.31	0.021	.1659534	.8621885	

- Interpretation of the coefficients: The odds ratio is for a one unit change in the predictor
- For this example 1.03 is the odds ratio for a one-year difference in age

Continuous explanatory variable

- When the explanatory variable is a continuous variable, there is no equivalent tabular method without dividing the variable into groups (e.g. age groups)
- You need to use logistic regression when you want to model the relationship between a dichotomous outcome and a continuous explanatory variable

Continuous explanatory variable

- To find the OR for a 10-year change in age

```
. logistic peth_audit_50 age_b, coef
```

```
Logistic regression                                Number of obs   =        324
                                                    LR chi2(1)      =         5.32
                                                    Prob > chi2     =        0.0210
Log likelihood = -221.81862                        Pseudo R2      =        0.0119
```

```
-----+-----
peth_audit_50 |      Coef.   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
          age_b |   .0295481   .0129978     2.27   0.023     .0040729     .0550234
          _cons |  -.972165   .4203565    -2.31   0.021    -1.796049    -.1482814
-----+-----
```

OR for a 10-year change in age = $\exp(10 \times 0.029) = 1.34$

Continuous explanatory variable

- To find the OR for a 10-year change in age

```
. lincom age_b*10
```

```
( 1) 10*[peth_audit_50]age_b = 0
```

peth_audi~50	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
-----+-----						
(1)	1.343773	.1746612	2.27	0.023	1.04157	1.733658

Continuous explanatory variable

- Another way to find the OR for a 10-year change in age

```
gen age_10 = age_b/10
```

```
. logistic peth_audit_50 age_10
```

```
Logistic regression
```

```
Number of obs   =      324  
LR chi2(1)      =       5.32  
Prob > chi2     =      0.0210  
Pseudo R2      =      0.0119
```

```
Log likelihood = -221.81862
```

	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
peth_audit_50						
age_10	1.343773	.1746612	2.27	0.023	1.04157	1.733658
_cons	.3782632	.1590054	-2.31	0.021	.1659534	.8621885

Checking logit-linearity

```
. tab agecat
```

agecat	Freq.	Percent	Cum.
15	80	24.69	24.69
25	136	41.98	66.67
35	80	24.69	91.36
45	28	8.64	100.00
Total	324	100.00	

```
logistic peth_audit_50 i.agecat
```

Logistic regression

Number of obs = 324
 LR chi2(3) = 8.22
 Prob > chi2 = 0.0416
 Pseudo R2 = 0.0183

Log likelihood = -220.3693

peth_audit_50	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
agecat						
25	1.525822	.4390615	1.47	0.142	.8680944	2.681887
35	2.254902	.728875	2.52	0.012	1.196699	4.248838
45	2.575758	1.16071	2.10	0.036	1.06495	6.229893
_cons	.6	.1385641	-2.21	0.027	.3815704	.9434694

Logistic regression with >1 explanatory variable

- In multiple logistic regression, the probability of success depends on more than one explanatory variable

$$p = \frac{e^{\alpha + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k}}{1 + e^{\alpha + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k}}$$

$$\text{and } \frac{p}{1-p} = e^{\alpha + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k}$$

- Therefore

$$\ln(p/(1-p)) = \alpha + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k$$

Interpretation of the coefficients

- Example: 2 independent variables, X_1 a numerical variable, X_2 a dichotomous exposure variable
- $\text{logit}(p) = \log(p/(1-p)) = \alpha + \beta_1 x_1 + \beta_2 x_2$
- When $x_2=1$ (exposed), holding x_1 constant,
$$\text{logit}(p_1) = \alpha + \beta_1 x_1 + \beta_2$$
- When $x_2=0$ (not exposed), holding x_1 constant,
$$\text{logit}(p_0) = \alpha + \beta_1 x_1$$

```
. logistic peth_audit_50 age_10 i.studyarm_n
```

Logistic regression

Number of obs = 324

LR chi2(2) = 6.00

Prob > chi2 = 0.0498

Log likelihood = -221.48068

Pseudo R2 = 0.0134

```
-----+-----  
peth_audit_50 | Odds Ratio   Std. Err.      z    P>|z|     [95% Conf. Interval]  
-----+-----  
      age_10 |    1.328877   .1737573     2.17   0.030     1.028457    1.717053  
  1.studyarm_n |    .8295213   .1886524    -0.82   0.411     .5311835    1.29542  
        _cons |    .4248297   .188248    -1.93   0.053     .1782524    1.012498  
-----+-----
```

0.83 is the odds ratio for PEth $\geq$ 50 for being in the minimally assessed vs. standard study arm when you hold age constant

To sum up...

OLS

- Response variable = Y
- Link function between the response variable and what is being modeled = 1
- Eq. $Y = a + bX$
- Distr. = Normal
- Var. = Constant

Logistic

- Response variable = Y
- Link function between the response variable and what is being modeled = logit
- $\ln(\pi/1-\pi) = a + bX$
- Distr. = Binomial
- Var = Not Constant

To sum up...

OLS

- $Y=a+bX$
- Interpretation of b
- Omnibus test =F
- Testing of b = t
- Goodness of fit= R^2

Logistic

- $\ln(\pi/1-\pi)=a+bX$
- Interpretation of b
- Omnibus test =-2LL
- Testing of b = Wald
- Goodness of fit= -2LL, pseudo- R^2

To sum up... assumptions

OLS

- Linearity
- Additivity
- Normal distribution
- Errors terms are independent

Logistic

- Non-Linearity ... however assumes a linear relationship between the logit of the dependent and the independent variables.
- Additivity
- Non-normal distribution
- Errors terms are independent

To sum up... assumptions

OLS

- OLS relies on the central limit theorem which states that 'as we increase the sample size of the statistics that make up the sampling distribution,..., the form of the distribution becomes increasingly symmetrical and bell-shaped

Logistic

- 'Maximum likelihood estimation (MLE) relies on large-sample asymptotic normality which means that reliability of estimates decline when there are few cases for each observed combination of X variables.'

- Next week – more multiple regression and multivariate logistic regression
- Review of methods this term
- Good website for Stata statistical methods:
<http://www.ats.ucla.edu/stat/stata/webbooks/>